

Asymptotic analysis of average case approximation complexity of Hilbert space valued random elements*

A. A. Khartov

October 17, 2014

Abstract

We study approximation properties of sequences of centered random elements X_d , $d \in \mathbb{N}$, with values in separable Hilbert spaces. We focus on sequences of tensor product-type and, in particular, degree-type random elements, which have covariance operators of corresponding tensor form. The average case approximation complexity $n^{X_d}(\varepsilon)$ is defined as the minimal number of continuous linear functionals that is needed to approximate X_d with relative 2-average error not exceeding a given threshold $\varepsilon \in (0, 1)$. In the paper we investigate $n^{X_d}(\varepsilon)$ for arbitrary fixed $\varepsilon \in (0, 1)$ and $d \rightarrow \infty$. Namely, we find criteria of (un)boundedness for $n^{X_d}(\varepsilon)$ on d and of tending $n^{X_d}(\varepsilon) \rightarrow \infty$, $d \rightarrow \infty$, for any fixed $\varepsilon \in (0, 1)$. In the latter case we obtain necessary and sufficient conditions for the following logarithmic asymptotics

$$\ln n^{X_d}(\varepsilon) = a_d + q(\varepsilon)b_d + o(b_d), \quad d \rightarrow \infty,$$

at continuity points of a non-decreasing function $q: (0, 1) \rightarrow \mathbb{R}$. Here $(a_d)_{d \in \mathbb{N}}$ is a sequence and $(b_d)_{d \in \mathbb{N}}$ is a positive sequence such that $b_d \rightarrow \infty$, $d \rightarrow \infty$. Under rather weak assumptions, we show that for tensor product-type random elements only special quantiles of self-decomposable or, in particular, stable (for tensor degrees) probability distributions appear as functions q in the asymptotics.

We apply our results to the tensor products of the Euler integrated processes with a given variation of smoothness parameters and to the tensor degrees of random elements with regularly varying eigenvalues of covariance operator.

1 Introduction

Let X be a centered random element of some normed space $(Q, \|\cdot\|_Q)$. Let us approximate X by the finite rank sums $\tilde{X}^{(n)} = \sum_{k=1}^n l_k(X)\psi_k$, where ψ_k are deterministic elements of Q and l_k are continuous linear functionals from the dual space Q^* . It is of theoretical and practical interest to make the relative average approximation error $(\mathbb{E}\|X - \tilde{X}^{(n)}\|_Q^2 / \mathbb{E}\|X\|_Q^2)^{1/2}$ smaller than a given error threshold ε by choosing appropriate $n \in \mathbb{N}$ and optimal ψ_k and l_k . Here we deal with *linear*

*The work was supported by the Government of the Russian Federation megagrant 11.G34.31.0026, by JSC “Gazprom Neft”, by the RFBR grant 13-01-00172, by the SPbGU grant 6.38.672.2013, and by the grant of Scientific school NSh-2504.2014.1.

approximation problem in average case setting (see [31]). We call the minimal suitable value of such n , denoted by $n^X(\varepsilon)$, the *average case approximation complexity* of the random element X (see [35]–[38]).

Now we consider a sequence of centered random elements X_d , $d \in \mathbb{N}$, with values in normed spaces $(Q_d, \|\cdot\|_{Q_d})$, $d \in \mathbb{N}$, respectively. When X_d , $d \in \mathbb{N}$, are somehow related, it is of interest to look at the behaviour of the quantity $n^{X_d}(\varepsilon)$, as d varies. In particular, for important *linear tensor product approximation problems* (see [27]–[29]) we have $Q_d = \otimes_{j=1}^d Q_{1,j}$ in appropriate sense, where $Q_{1,j}$, $j \in \mathbb{N}$, are some normed spaces. Here the covariances operators K^{X_d} of X_d , $d \in \mathbb{N}$, also have the appropriate tensor product form $K^{X_d} = \otimes_{j=1}^d K^{X_{1,j}}$, $d \in \mathbb{N}$, where $K^{X_{1,j}}$ is a covariance operator of given $Q_{1,j}$ -valued centered random element $X_{1,j}$, $j \in \mathbb{N}$. Such X_d is called the *tensor product of $X_{1,1}, \dots, X_{1,d}$* . If all $Q_{1,j}$ in Q_d are the same and all $K^{X_{1,j}}$ in K^{X_d} are equal, then such X_d is said to be the *tensor degree of $X_{1,1}$* . A classical example of such objects is the well known *d-parametric Brownian sheet* (or the *Wiener–Chentsov random field*, see [1] and [15]). It, being considered as a random element of the space $L_2([0, 1]^d)$, is a tensor degree of the Wiener process as a random element of $L_2([0, 1])$.

The described multivariate approximation problems find many applications in simulation methods, statistics, physics, and computational finance (see [26]). In this connection, tractability questions becomes rather actual now. A sequence of approximation problems for X_d , $d \in \mathbb{N}$ is *weakly tractable* if $n^{X_d}(\varepsilon)$ is not exponential in d or/and ε^{-1} . Otherwise, the sequence of the problems is *intractable*. Special subclasses of weakly tractable problems are distinguished depending on the types of majorants for the quantity $n^{X_d}(\varepsilon)$ for all $d \in \mathbb{N}$ and $\varepsilon \in (0, 1)$. There exist some results concerning the tractability of the described linear (tensor product) approximation problems, in which Q_d are separable Hilbert spaces (see [16], [17] and [27]–[29]). We will consider these problems within other less explored setting, namely, when ε is arbitrarily close to zero but *fixed* and d goes to infinity. As noted in the book [27] (see p. 6 and 289), such setting, being more appropriate for some models in computational finance, is also important. But the author is aware, in fact, of only one article [18] by M. A. Lifshits and E. V. Tulyakova on this subject. Our purpose is to complement their results.

The paper is organized as follows. In Section 2 we provide a formal problem setting. In Section 3 we consider sequences of random elements X_d , $d \in \mathbb{N}$, of separable Hilbert spaces. In particular, we find necessary and sufficient conditions that for almost all $\varepsilon \in (0, 1)$ the quantity $n^{X_d}(\varepsilon)$ has a special form of logarithmic asymptotics as $d \rightarrow \infty$. In Sections 4 and 5 the same problem is solved for the sequences of tensor product-type and degree-type random elements, respectively. We show that, under rather weak assumption, $n^{X_d}(\varepsilon)$ depends on ε according to some *self-decomposable* or, in particular, *stable probability distribution*. We apply the obtained criteria to tensor products of the Euler integrated processes with a given variation of the smoothness parameters and to tensor degrees of random elements with a given regular variation of eigenvalues of covariance operator. In Appendix we provide necessary facts from probability theory about self-decomposable and stable distributions and also about related limit theorems, which are the main tools of our work.

Throughout the article, we use the following notation. We write $a_n \sim b_n$ iff $a_n/b_n \rightarrow 1$, $n \rightarrow \infty$. We denote by \mathbb{N} and \mathbb{R} the sets of positive integer and real numbers, respectively. We set $\ln_+ x := \max\{1, \ln x\}$ for all $x > 0$. The quantity $\mathbb{1}(A)$ equals one for the true logic propositions A and zero for the false ones. We always use $\|\cdot\|_B$ for the norm, which some space B is equipped with. For any function f we will denote by $\mathbf{C}(f)$ the set of all its continuity points and by f^{-1} the generalized inverse function $f^{-1}(y) := \inf\{x \in \mathbb{R} : f(x) \geq y\}$, where y is from the range of f . By *distribution function* F we mean the non-decreasing function F on \mathbb{R} that is right-continuous on \mathbb{R} , $\lim_{x \rightarrow -\infty} F(x) = 0$, and

$\lim_{x \rightarrow \infty} F(x) = 1$. Following [21], the boundaries of growth points of distribution function F will be denoted by

$$\text{left } F := \inf\{x \in \mathbb{R} : F(x) > 0\} \geq -\infty, \quad \text{right } F := \sup\{x \in \mathbb{R} : F(x) < 1\} \leq \infty.$$

A distribution function F is called *degenerate* if $F(x) = \mathbb{1}(x \geq a)$ for any x and some constant a .

2 Basic definitions and problem setting

We consider a sequence of random elements X_d , $d \in \mathbb{N}$, with values in separable Hilbert spaces H_d , $d \in \mathbb{N}$, respectively. Assume that every X_d has zero mean and $\mathbb{E} \|X_d\|_{H_d}^2 < \infty$, $d \in \mathbb{N}$. We will investigate the *average case approximation complexity* (simply the *approximation complexity* for short) of X_d , $d \in \mathbb{N}$:

$$n^{X_d}(\varepsilon) := \min\{n \in \mathbb{N} : e^{X_d}(n) \leq \varepsilon e^{X_d}(0)\}, \quad (1)$$

where $\varepsilon \in (0, 1)$ is a given error threshold, and

$$e^{X_d}(n) := \inf\left\{\left(\mathbb{E} \|X_d - \tilde{X}_d^{(n)}\|_{H_d}^2\right)^{1/2} : \tilde{X}_d^{(n)} \in \mathcal{A}_n^{X_d}\right\}$$

is the smallest 2-average error among all linear approximations of X_d , $d \in \mathbb{N}$, having rank $n \in \mathbb{N}$. The corresponding classes of linear algorithms are denoted by

$$\mathcal{A}_n^{X_d} := \left\{\sum_{m=1}^n l_m(X_d) \psi_m : \psi_m \in H_d, l_m \in H_d^*\right\}, \quad d \in \mathbb{N}, \quad n \in \mathbb{N}.$$

We work with *relative errors*, thus taking into account the following “size” of X_d :

$$e^{X_d}(0) := \left(\mathbb{E} \|X_d\|_{H_d}^2\right)^{1/2},$$

which is the approximation error of X_d by zero element of H_d . The approximation complexity $n^{X_d}(\varepsilon)$ is considered as a function depending on two variables d and ε . The general goal is to understand the character of this dependence for the given sequence $(X_d)_{d \in \mathbb{N}}$.

The linear tensor product approximation problems are of our particular interest. We will study these only within the following construction. We suppose that every H_d is defined by the Hilbertian tensor product $H_d := \otimes_{j=1}^d H_{1,j}$, where every $H_{1,j}$ is a given separable Hilbert space, $j \in \mathbb{N}$. We also suppose that for every d the random element X_d is a *tensor product* of some zero-mean $H_{1,j}$ -valued random elements $X_{1,j}$, $j = 1, \dots, d$, i.e. X_d has zero mean and the covariance operator $K^{X_d} := \otimes_{j=1}^d K^{X_{1,j}}$, where $K^{X_{1,j}}$ is a covariance operator of $X_{1,j}$ for every $j \in \mathbb{N}$. Following [14] and [19], for such X_d we will write $X_d = \otimes_{j=1}^d X_{1,j}$ for short. In particular, if $H_{1,j}$ are the same and $K^{X_{1,1}} = \dots = K^{X_{1,d}}$, then the element X_d is called the *tensor degree* of $X_{1,1}$ and we will write $X_d = X_{1,1}^{\otimes d}$ in this case.

We now turn to tractability of approximation problems. Let us give basic definitions related to the tractability in context of our setting (1). Let APP denote a sequence of linear approximation problems for X_d , $d \in \mathbb{N}$. We say that

- APP is *weakly tractable* iff

$$\lim_{d+\varepsilon^{-1} \rightarrow \infty} \frac{\ln_+ n^{X_d}(\varepsilon)}{d + \varepsilon^{-1}} = 0; \quad (2)$$

- APP is *quasi-polynomially tractable* iff there are numbers $C > 0$ and $s \geq 0$ such that

$$n^{X_d}(\varepsilon) \leq C \exp\{s(1 + \ln \varepsilon^{-1}) \ln_+ d\} \quad \text{for all } d \in \mathbb{N}, \varepsilon \in (0, 1); \quad (3)$$

- APP is *polynomially tractable* iff there are numbers $C > 0$, $s \geq 0$, and $p \geq 0$ such that

$$n^{X_d}(\varepsilon) \leq C \varepsilon^{-s} d^p \quad \text{for all } d \in \mathbb{N}, \varepsilon \in (0, 1); \quad (4)$$

- APP is *strong polynomially tractable* iff there are numbers $C > 0$ and $s \geq 0$ such that

$$n^{X_d}(\varepsilon) \leq C \varepsilon^{-s} \quad \text{for all } d \in \mathbb{N}, \varepsilon \in (0, 1). \quad (5)$$

If APP is not weakly tractable, then it is called *intractable*. In particular, if $n^{X_d}(\varepsilon)$ increases at least exponentially in d , then we say that APP has a *curse of dimensionality*.

Let us give a short review of some results concerning the tractability in our setting. For the linear approximation problems there exist criteria of each type of tractability in terms of eigenvalues of the operators K^{X_d} , $d \in \mathbb{N}$. These results can be found in [27] (see p. 245, 256). The described linear tensor product approximation problems were investigated in the recent paper [16], where the necessary and sufficient conditions for all types of tractability were given in terms of eigenvalues of the covariance operators $K^{X_{1,j}}$, $j \in \mathbb{N}$. Also the tensor product-type Korobov kernels were studied there. In [17] the authors obtained results concerning tractability of tensor products of the Euler and Wiener integrated processes with varying smoothness parameters.

In the paper we investigate the approximation complexity $n^{X_d}(\varepsilon)$ of general and tensor product-type random elements for arbitrary fixed $\varepsilon \in (0, 1)$ and $d \rightarrow \infty$. Namely, we search necessary and sufficient conditions that $n^{X_d}(\varepsilon)$ has the following logarithmic asymptotics

$$\ln n^{X_d}(\varepsilon) = a_d + q(\varepsilon)b_d + o(b_d), \quad d \rightarrow \infty, \quad \text{for all } \varepsilon \in \mathbf{C}(q), \quad (6)$$

where $(a_d)_{d \in \mathbb{N}}$ is a sequence, $(b_d)_{d \in \mathbb{N}}$ is a positive sequence such that $b_d \rightarrow \infty$, $d \rightarrow \infty$, and the function $q: (0, 1) \rightarrow \mathbb{R}$ is non-increasing. To consider the asymptotics of the form (6) is quite natural and, moreover, as we will see below, it is inherent in the linear tensor product approximation problems. The author is not aware of any results, which could provide the full solution in such setting. But there exist three papers [18]–[20] concerning asymptotic analysis of $n^{X_d}(\varepsilon)$ for fixed ε and $d \rightarrow \infty$, where only the first one deals with objects of our interest. In [18] the authors considered only the tensor degree-type random elements under some assumptions. We will discuss the corresponding theorem in Section 5.

As in case of tractability, eigenvalues of covariance operators of X_d , $d \in \mathbb{N}$ or $X_{1,j}$, $j \in \mathbb{N}$, play a crucial role in asymptotic analysis of the quantity $n^{X_d}(\varepsilon)$ as $d \rightarrow \infty$. For convenience, throughout the paper we will use the following unified notation for the correlation characteristics. For a given Hilbertian random element Z we will denote by K^Z its covariance operator. The sequences $(\lambda_k^Z)_{k \in \mathbb{N}}$ and $(\psi_k^Z)_{k \in \mathbb{N}}$ will denote the non-increasing sequence of eigenvalues and the corresponding sequence of eigenvectors of K^Z , respectively, i.e. $K^Z \psi_k^Z = \lambda_k^Z \psi_k^Z$, $k \in \mathbb{N}$. If Z is a random element of p -dimensional space, then we formally set $\lambda_k^Z := 0$, and $\psi_k^Z := 0$ for $k > p$. The trace of K^Z will be denoted by Λ^Z , thus $\Lambda^Z = \sum_{k=1}^{\infty} \lambda_k^Z$. We will also use the notation $\bar{\lambda}_k^Z := \lambda_k^Z / \Lambda^Z$, $k \in \mathbb{N}$, (i.e. $\sum_{k=1}^{\infty} \bar{\lambda}_k^Z = 1$) and $\mathbf{m}^Z(x) := \sum_{k=1}^{\infty} \mathbf{1}(\bar{\lambda}_k^Z = x)$ for any $x > 0$.

3 Approximation of general random elements

Here we consider a general sequence of random elements X_d , $d \in \mathbb{N}$, of abstract separable Hilbert spaces H_d , $d \in \mathbb{N}$, respectively, without any assumptions on the spectral structure of the corresponding covariance operators K^{X_d} , $d \in \mathbb{N}$. Let H_d be equipped with inner product $(\cdot, \cdot)_{H_d}$. We always assume that every X_d has zero mean and satisfies $\mathbb{E} \|X_d\|_{H_d}^2 < \infty$, $d \in \mathbb{N}$. Then self-adjoint non-negative definite operators K^{X_d} , $d \in \mathbb{N}$, have the finite traces:

$$\Lambda^{X_d} = \sum_{k=1}^{\infty} \lambda_k^{X_d} = \mathbb{E} \|X_d\|_{H_d}^2 = e^{X_d}(0)^2 < \infty, \quad d \in \mathbb{N}. \quad (7)$$

To omit the pathological cases, we always assume that $\lambda_1^{X_d} > 0$ for all $d \in \mathbb{N}$.

It is well known (see [40]) that for any $n \in \mathbb{N}$ the following random element

$$\tilde{X}_d^{(n)} := \sum_{k=1}^n (X_d, \psi_k^{X_d})_{H_d} \psi_k^{X_d} \in \mathcal{A}_n^{X_d} \quad (8)$$

minimizes the 2-average case error. Hence formula (1) is reduced to

$$n^{X_d}(\varepsilon) = \min \left\{ n \in \mathbb{N} : \mathbb{E} \|X_d - \tilde{X}_d^{(n)}\|_{H_d}^2 \leq \varepsilon^2 \mathbb{E} \|X_d\|_{H_d}^2 \right\}.$$

From (7) and (8) we infer the following representation of the approximation complexity

$$n^{X_d}(\varepsilon) = \min \left\{ n \in \mathbb{N} : \sum_{k=n+1}^{\infty} \lambda_k^{X_d} \leq \varepsilon^2 \Lambda^{X_d} \right\}.$$

In terms of $\bar{\lambda}_k^{X_d}$, $k \in \mathbb{N}$, it takes the form

$$n^{X_d}(\varepsilon) = \min \left\{ n \in \mathbb{N} : \sum_{k=n+1}^{\infty} \bar{\lambda}_k^{X_d} \leq \varepsilon^2 \right\} \quad (9)$$

$$= \min \left\{ n \in \mathbb{N} : \sum_{k=1}^n \bar{\lambda}_k^{X_d} \geq 1 - \varepsilon^2 \right\}. \quad (10)$$

3.1 Boundedness of the approximation complexity

Before proceeding to the asymptotic analysis of the quantity $n^{X_d}(\varepsilon)$, we find criteria of its boundedness and unboundedness on d for any fixed $\varepsilon \in (0, 1)$. The next simple proposition provides the conditions for $n^{X_d}(\varepsilon) \rightarrow \infty$ as $d \rightarrow \infty$ in terms of the first normed eigenvalues $\bar{\lambda}_1^{X_d}$, $d \in \mathbb{N}$.

Proposition 1 *The following conditions are equivalent:*

- (i) $\lim_{d \rightarrow \infty} n^{X_d}(\varepsilon) = \infty$ for all $\varepsilon \in (0, 1)$;
- (ii) $\lim_{d \rightarrow \infty} \bar{\lambda}_1^{X_d} = 0$.

Proof of Proposition 1. (i) \Rightarrow (ii). Suppose that, contrary to our claim, there exists a subsequence $(d_l)_{l \in \mathbb{N}}$ such that $c := \lim_{l \rightarrow \infty} \bar{\lambda}_1^{X_{d_l}} > 0$. Choose $\varepsilon \in (0, 1)$ to satisfy $1 - \varepsilon^2 < c$. Then, according to (10), we have $n^{X_{d_l}}(\varepsilon) = 1$ for all sufficiently large $l \in \mathbb{N}$. This contradicts our assumption (i).

(i) \Leftarrow (ii). For all $\varepsilon \in (0, 1)$ the statement $n^{X_d}(\varepsilon) \rightarrow \infty$, $d \rightarrow \infty$, follows from the formula (10) and the inequality:

$$n^{X_d}(\varepsilon) \geq \sum_{k=1}^{n^{X_d}(\varepsilon)} \frac{\bar{\lambda}_k^{X_d}}{\bar{\lambda}_1^{X_d}} \geq \frac{1 - \varepsilon^2}{\bar{\lambda}_1^{X_d}}. \quad \square \quad (11)$$

As a consequence of the previous proposition, we obtain the following criterion of unboundedness of $n^{X_d}(\varepsilon)$ on $d \in \mathbb{N}$ for any fixed $\varepsilon \in (0, 1)$.

Proposition 2 *The following conditions are equivalent:*

- (i) $\sup_{d \in \mathbb{N}} n^{X_d}(\varepsilon) = \infty$ for all $\varepsilon \in (0, 1)$;
- (ii) $\inf_{d \in \mathbb{N}} \bar{\lambda}_1^{X_d} = 0$.

Let us formulate the criterion of boundedness of the approximation complexity $n^{X_d}(\varepsilon)$ on $d \in \mathbb{N}$ for any fixed $\varepsilon \in (0, 1)$.

Proposition 3 *The following conditions are equivalent:*

- (i) $\sup_{d \in \mathbb{N}} n^{X_d}(\varepsilon) < \infty$ for all $\varepsilon \in (0, 1)$;
- (ii) $\lim_{t \rightarrow 0} \sup_{d \in \mathbb{N}} \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_d} \mathbb{1}(\bar{\lambda}_k^{X_d} < t) = 0$.

Before proving this proposition, we introduce the following auxiliary quantity

$$\bar{\lambda}^{X_d}(\varepsilon) := \bar{\lambda}_{n^{X_d}(\varepsilon)}^{X_d}, \quad d \in \mathbb{N}, \quad \varepsilon \in (0, 1), \quad (12)$$

which admits the following representation

$$\bar{\lambda}^{X_d}(\varepsilon) = \sup \left\{ x \in \mathbb{R} : \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_d} \mathbb{1}(\bar{\lambda}_k^{X_d} < x) \leq \varepsilon^2 \right\}. \quad (13)$$

Obtain useful inequalities connecting $\bar{\lambda}^{X_d}(\varepsilon)$ and $n^{X_d}(\varepsilon)$. First, from the definitions of these quantities we conclude that

$$n^{X_d}(\varepsilon) \leq \bar{\lambda}^{X_d}(\varepsilon)^{-1} \quad \text{for all } d \in \mathbb{N}, \quad \varepsilon \in (0, 1). \quad (14)$$

Next, for any $\varepsilon_1 \in (0, 1)$, $\varepsilon_2 \in (\varepsilon_1, 1)$, and $d \in \mathbb{N}$ we have

$$n^{X_d}(\varepsilon_1) \geq n^{X_d}(\varepsilon_1) - n^{X_d}(\varepsilon_2) + 1 \geq \sum_{k=n^{X_d}(\varepsilon_2)+1}^{n^{X_d}(\varepsilon_1)} \frac{\bar{\lambda}_k^{X_d}}{\bar{\lambda}^{X_d}(\varepsilon_2)} + 1,$$

where

$$\begin{aligned}
\sum_{k=n^{X_d(\varepsilon_2)}+1}^{n^{X_d(\varepsilon_1)}} \bar{\lambda}_k^{X_d} &= \sum_{k=1}^{n^{X_d(\varepsilon_1)}} \bar{\lambda}_k^{X_d} - \sum_{k=1}^{n^{X_d(\varepsilon_2)}} \bar{\lambda}_k^{X_d} \\
&\geq (1 - \varepsilon_1^2) - ((1 - \varepsilon_2^2) + \bar{\lambda}^{X_d}(\varepsilon_2)) \\
&= \varepsilon_2^2 - \varepsilon_1^2 - \bar{\lambda}^{X_d}(\varepsilon_2).
\end{aligned}$$

Finally, we get

$$n^{X_d}(\varepsilon_1) \geq \frac{\varepsilon_2^2 - \varepsilon_1^2}{\bar{\lambda}^{X_d}(\varepsilon_2)} \quad \text{for all } d \in \mathbb{N}, \varepsilon_1 \in (0, 1), \varepsilon_2 \in (\varepsilon_1, 1). \quad (15)$$

Proof of Proposition 3. First, we show that the condition (i) is equivalent to

$$\inf_{d \in \mathbb{N}} \bar{\lambda}^{X_d}(\varepsilon) > 0 \quad \text{for all } \varepsilon \in (0, 1). \quad (16)$$

Indeed, from (14) we conclude the implication (16) \Rightarrow (i). Next, let us fix any $\varepsilon \in (0, 1)$ and $c \in (0, 1)$. Using inequality (15) with $\varepsilon_1 = c\varepsilon$ and $\varepsilon_2 = \varepsilon$, we obtain

$$\bar{\lambda}^{X_d}(\varepsilon) \geq \frac{(1 - c^2)\varepsilon^2}{\sup_{d \in \mathbb{N}} n^{X_d}(c\varepsilon)} \quad \text{for all } d \in \mathbb{N}.$$

From this we get the implication (i) \Rightarrow (16).

It only remains to verify (ii) \Leftrightarrow (16). But it follows from the assertion that for any $\varepsilon \in (0, 1)$ and some $t_\varepsilon \in (0, 1)$ we have $\inf_{d \in \mathbb{N}} \bar{\lambda}^{X_d}(\varepsilon) \geq t_\varepsilon$ iff $\sup_{d \in \mathbb{N}} \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_d} \mathbf{1}(\bar{\lambda}_k^{X_d} < t_\varepsilon) \leq \varepsilon^2$. \square

3.2 Logarithmic asymptotics of the approximation complexity

Here we start asymptotic analysis of the approximation complexity $n^{X_d}(\varepsilon)$. Our next theorem establishes a connection between the asymptotics of the form (6), for given $(a_d)_{d \in \mathbb{N}}$ and $(b_d)_{d \in \mathbb{N}}$, and the convergence of the following distribution functions

$$G_{a_d, b_d}^{X_d}(x) := \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_d} \mathbf{1}(\bar{\lambda}_k^{X_d} \geq e^{-a_d - x b_d}), \quad x \in \mathbb{R}, \quad d \in \mathbb{N}. \quad (17)$$

Theorem 1 *Let $(a_d)_{d \in \mathbb{N}}$ be a sequence, $(b_d)_{d \in \mathbb{N}}$ be a positive sequence such that $b_d \rightarrow \infty$, $d \rightarrow \infty$. Let a non-increasing function $q: (0, 1) \rightarrow \mathbb{R}$ and a distribution function G satisfy the equation $q(\varepsilon) = G^{-1}(1 - \varepsilon^2)$ for all $\varepsilon \in \mathbf{C}(q)$. For the asymptotics*

$$\ln n^{X_d}(\varepsilon) = a_d + q(\varepsilon)b_d + o(b_d), \quad d \rightarrow \infty, \quad \text{for all } \varepsilon \in \mathbf{C}(q), \quad (18)$$

it is necessary and sufficient to have the weak convergence¹ $G_{a_d, b_d}^{X_d} \Rightarrow G$ as $d \rightarrow \infty$, i.e.

$$\lim_{d \rightarrow \infty} G_{a_d, b_d}^{X_d}(x) = G(x) \quad \text{for all } x \in \mathbf{C}(G). \quad (19)$$

¹We follow the definition from [30], p. 16. Equivalent definitions can be found in [6], p. 248–251.

Proof of Theorem 1. It is well known that condition (19) is equivalent to the convergence $\lim_{d \rightarrow \infty} (G_{a_d, b_d}^{X_d})^{-1}(y) = G^{-1}(y)$ for all continuity points y of G^{-1} (see [39], p. 305). By theorem of continuity of composite functions, we have $\varepsilon \in \mathbf{C}(q) \Leftrightarrow 1 - \varepsilon^2 \in \mathbf{C}(G^{-1})$. Therefore (19) is equivalent to the following condition

$$\lim_{d \rightarrow \infty} (G_{a_d, b_d}^{X_d})^{-1}(1 - \varepsilon^2) = q(\varepsilon) \quad \text{for all } \varepsilon \in \mathbf{C}(q).$$

According to (12) and (13), we note that

$$\begin{aligned} |\ln \bar{\lambda}^{X_d}(\varepsilon)| &= \inf \left\{ y \in \mathbb{R} : \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_d} \mathbf{1}(\bar{\lambda}_k^{X_d} \geq e^{-y}) \geq 1 - \varepsilon^2 \right\} \\ &= \inf \left\{ y \in \mathbb{R} : G_{a_d, b_d}^{X_d}((y - a_d)/b_d) \geq 1 - \varepsilon^2 \right\} \\ &= a_d + b_d \cdot \inf \left\{ z \in \mathbb{R} : G_{a_d, b_d}^{X_d}(z) \geq 1 - \varepsilon^2 \right\} \\ &= a_d + b_d \cdot (G_{a_d, b_d}^{X_d})^{-1}(1 - \varepsilon^2). \end{aligned} \tag{20}$$

Thus (19) is equivalent to the following condition

$$|\ln \bar{\lambda}^{X_d}(\varepsilon)| = a_d + q(\varepsilon)b_d + o(b_d), \quad d \rightarrow \infty, \quad \text{for all } \varepsilon \in \mathbf{C}(q). \tag{21}$$

Let us prove the implication (21) \Rightarrow (18). Fix $\varepsilon \in \mathbf{C}(q)$ and arbitrarily small $h > 0$. From (21) and (14) we have the inequality $\ln n^{X_d}(\varepsilon) \leq a_d + q(\varepsilon)b_d + hb_d$ for all large enough $d \in \mathbb{N}$. Since the function q is non-increasing, the set $\mathbf{C}(q)$ is dense in the interval $(0, 1)$. There exists $\tau_1 \in (1, 1/\varepsilon)$ such that $\tau_1\varepsilon \in \mathbf{C}(q)$ and $q(\tau_1\varepsilon) - q(\varepsilon) \geq -h$. By the inequality (15), we have

$$\ln n^{X_d}(\varepsilon) \geq |\ln \bar{\lambda}^{X_d}(\tau_1\varepsilon)| + \ln((\tau_1\varepsilon)^2 - \varepsilon^2).$$

According to (21), for all large enough $d \in \mathbb{N}$ we obtain

$$\begin{aligned} \ln n^{X_d}(\varepsilon) &\geq a_d + q(\tau_1\varepsilon)b_d - hb_d + \ln((\tau_1\varepsilon)^2 - \varepsilon^2) \\ &\geq a_d + q(\varepsilon)b_d - 2hb_d + \ln((\tau_1\varepsilon)^2 - \varepsilon^2). \end{aligned}$$

Since $b_d \rightarrow \infty$ as $d \rightarrow \infty$, it follows that $\ln n^{X_d}(\varepsilon) \geq a_d + q(\varepsilon)b_d - 3hb_d$ for all large enough $d \in \mathbb{N}$. Thus the asymptotics (18) follows from the obtained estimates for $n^{X_d}(\varepsilon)$.

Prove (18) \Rightarrow (21). Fix $\varepsilon \in \mathbf{C}(q)$ and arbitrarily small $h > 0$. From (18) and (14) we obtain the inequality $|\ln \bar{\lambda}^{X_d}(\varepsilon)| \geq a_d + q(\varepsilon)b_d - hb_d$ for all large enough $d \in \mathbb{N}$. Find $\tau_2 \in (0, 1)$ such that $\tau_2\varepsilon \in \mathbf{C}(q)$ and $q(\tau_2\varepsilon) - q(\varepsilon) \leq h$. The inequality (15) yields

$$|\ln \bar{\lambda}^{X_d}(\varepsilon)| \leq \ln n^{X_d}(\tau_2\varepsilon) - \ln(\varepsilon^2 - (\tau_2\varepsilon)^2).$$

On account of the asymptotics (18), for all large enough $d \in \mathbb{N}$ we obtain

$$\begin{aligned} |\ln \bar{\lambda}^{X_d}(\varepsilon)| &\leq a_d + q(\tau_2\varepsilon)b_d + hb_d - \ln(\varepsilon^2 - (\tau_2\varepsilon)^2) \\ &\leq a_d + q(\varepsilon)b_d + 2hb_d - \ln(\varepsilon^2 - (\tau_2\varepsilon)^2). \end{aligned}$$

Since $b_d \rightarrow \infty$ as $d \rightarrow \infty$, we have $|\ln \bar{\lambda}^{X_d}(\varepsilon)| \leq a_d + q(\varepsilon)b_d + 3hb_d$ for all large enough $d \in \mathbb{N}$. The obtained estimates yield the required asymptotics (21). \square

Let us mention some elementary facts for the functions G and q from Theorem 1.

Remark 1 *If for any given $\varepsilon \in (0, 1)$ the distribution function G strongly increases on the right (left) neighbourhood of $q(\varepsilon)$, then q is left-(right-)continuous at ε .*

Indeed, suppose that G strongly increases on the right neighbourhood of $q(\varepsilon)$ (the left case is similar). Then for any $\delta > 0$ we have $G(q(\varepsilon) + \delta) > 1 - \varepsilon^2$ and, consequently, there exists $\tau_\delta > 0$ such that $G(q(\varepsilon) + \delta) > 1 - (\varepsilon - \tau)^2$ for all $\tau \in (0, \tau_\delta)$. Applying G^{-1} to the previous inequality, we obtain $q(\varepsilon - \tau) - q(\varepsilon) \leq \delta$, which gives left-continuity of q at ε by monotony of q .

As a consequence of this remark, we provide the following useful note.

Remark 2 *If the distribution function G is degenerate or strongly increases on the non-empty interval $(\text{lxt } G, \text{rxt } G)$, then q is continuous on $(0, 1)$.*

4 Approximation of tensor product-type random elements

In this section we consider sequences of tensor product-type random elements. Suppose that we have a sequence of zero-mean random elements $X_{1,j}$, $j \in \mathbb{N}$, of separable Hilbert spaces $H_{1,j}$, $j \in \mathbb{N}$, respectively. Assume that every $X_{1,j}$ satisfies $\mathbb{E} \|X_{1,j}\|_{H_{1,j}}^2 < \infty$, $j \in \mathbb{N}$. Let $X_d = \otimes_{j=1}^d X_{1,j}$, $d \in \mathbb{N}$. It is well known that eigenvalues and eigenvectors of covariance operator K^{X_d} have the following multiplicative form:

$$\prod_{j=1}^d \lambda_{k_j}^{X_{1,j}}, \quad \bigotimes_{j=1}^d \psi_{k_j}^{X_{1,j}}, \quad k_1, k_2, \dots, k_d \in \mathbb{N}. \quad (22)$$

Hence for traces Λ^{X_d} of K^{X_d} , $d \in \mathbb{N}$, we have the formula

$$\Lambda^{X_d} = \sum_{k_1, k_2, \dots, k_d \in \mathbb{N}} \prod_{j=1}^d \lambda_{k_j}^{X_{1,j}} = \prod_{j=1}^d \sum_{i=1}^{\infty} \lambda_i^{X_{1,j}} = \prod_{j=1}^d \Lambda^{X_{1,j}}. \quad (23)$$

Throughout this section we assume that $\lambda_1^{X_{1,j}} > 0$ for all $j \in \mathbb{N}$. Since the first (the maximal) eigenvalue $\lambda_1^{X_d}$ of K^{X_d} equals $\prod_{j=1}^d \lambda_1^{X_{1,j}}$, this convention is equivalent to the assumption from Section 3, namely $\lambda_1^{X_d} > 0$, $d \in \mathbb{N}$.

4.1 Boundedness of the approximation complexity

By analogy with Section 3, we first consider the boundedness conditions of the approximation complexity $n^{X_d}(\varepsilon)$ on d for any fixed $\varepsilon \in (0, 1)$. For described tensor product-type random elements X_d , $d \in \mathbb{N}$, the following propositions show that for any fixed $\varepsilon \in (0, 1)$ either the quantity $n^{X_d}(\varepsilon)$ tends to infinity as $d \rightarrow \infty$ or it is a bounded function on $d \in \mathbb{N}$.

Proposition 4 *The following conditions are equivalent:*

$$(i) \quad \lim_{d \rightarrow \infty} n^{X_d}(\varepsilon) = \infty \quad \text{for all } \varepsilon \in (0, 1);$$

$$(ii) \quad \sum_{j=1}^{\infty} \sum_{k=2}^{\infty} \frac{\lambda_k^{X_{1,j}}}{\lambda_1^{X_{1,j}}} = \infty.$$

Proof of Proposition 4. By Proposition 1, the relation (i) is equivalent to the convergence $\lim_{d \rightarrow \infty} (\Lambda^{X_d} / \lambda_1^{X_d}) = \infty$, where, as it is easily seen:

$$\frac{\Lambda^{X_d}}{\lambda_1^{X_d}} = \prod_{j=1}^d \frac{\Lambda^{X_{1,j}}}{\lambda_1^{X_{1,j}}} = \prod_{j=1}^d \left(1 + \sum_{k=2}^{\infty} \frac{\lambda_k^{X_{1,j}}}{\lambda_1^{X_{1,j}}} \right). \quad (24)$$

The last product goes to infinity as $d \rightarrow \infty$ iff the relation (ii) holds. \square

Proposition 5 *The following conditions are equivalent:*

$$(i) \quad \sup_{d \in \mathbb{N}} n^{X_d}(\varepsilon) < \infty \quad \text{for all} \quad \varepsilon \in (0, 1);$$

$$(ii) \quad \sum_{j=1}^{\infty} \sum_{k=2}^{\infty} \frac{\lambda_k^{X_{1,j}}}{\lambda_1^{X_{1,j}}} < \infty.$$

Proof of Proposition 5. First, we note that (ii) is equivalent to

$$C := \sup_{d \in \mathbb{N}} (\Lambda^{X_d} / \lambda_1^{X_d}) < \infty.$$

Under the condition (i), the last directly follows from Proposition 2. It only remains to check that the assumption $C < \infty$ implies (i). Let $\hat{\lambda}_k^{X_\infty}$, $k \in \mathbb{N}$ be renumbered sequence of numbers $\prod_{j=1}^{\infty} (\lambda_{k_j}^{X_{1,j}} / \lambda_1^{X_{1,j}})$, where $k_j \in \mathbb{N}$ and $\lim_{j \rightarrow \infty} k_j = 1$. According to the multiplicative structure of $(\lambda_k^{X_d})_{k \in \mathbb{N}}$ (see (22)), we have

$$\begin{aligned} \sup_{d \in \mathbb{N}} \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_d} \mathbb{1}(\bar{\lambda}_k^{X_d} < t) &\leq \sup_{d \in \mathbb{N}} \sum_{k=1}^{\infty} (\lambda_k^{X_d} / \lambda_1^{X_d}) \mathbb{1}(\lambda_k^{X_d} / \lambda_1^{X_d} < Ct) \\ &= \sup_{d \in \mathbb{N}} \sum_{k \in \mathbb{N}^d} \prod_{j=1}^d (\lambda_{k_j}^{X_{1,j}} / \lambda_1^{X_{1,j}}) \mathbb{1} \left(\prod_{j=1}^d (\lambda_{k_j}^{X_{1,j}} / \lambda_1^{X_{1,j}}) < Ct \right) \\ &\leq \sum_{k=1}^{\infty} \hat{\lambda}_k^{X_\infty} \mathbb{1}(\hat{\lambda}_k^{X_\infty} < Ct). \end{aligned}$$

Since, by representation (24), $C = \sum_{k=1}^{\infty} \hat{\lambda}_k^{X_\infty} < \infty$, we obtain the equality

$$\lim_{t \rightarrow 0} \sup_{d \in \mathbb{N}} \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_d} \mathbb{1}(\bar{\lambda}_k^{X_d} < t) = \lim_{t \rightarrow 0} \sum_{k=1}^{\infty} \hat{\lambda}_k^{X_\infty} \mathbb{1}(\hat{\lambda}_k^{X_\infty} < Ct) = 0,$$

which is sufficient for (i) by Proposition 3. \square

4.2 Logarithmic asymptotics of the approximation complexity

Here we will obtain criteria for the asymptotics (6) under the following assumption

$$\lim_{d \rightarrow \infty} \max_{j=1, \dots, d} \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} < e^{-\tau b_d}) = 0 \quad \text{for all} \quad \tau > 0, \quad (25)$$

which is rather weak, because of the next assertion.

Remark 3 The condition $\lim_{j \rightarrow \infty} \bar{\lambda}_1^{X_{1,j}} = 1$ is sufficient for (25) for any sequence $(b_d)_{d \in \mathbb{N}}$ such that $b_d \rightarrow \infty$, $d \rightarrow \infty$.

Indeed, let us fix $\tau > 0$ and choose $j_\delta \in \mathbb{N}$ such that

$$1 - \bar{\lambda}_1^{X_{1,j}} = \sum_{k=2}^{\infty} \bar{\lambda}_k^{X_{1,j}} \leq \delta \quad \text{for all } j \geq j_\delta. \quad (26)$$

For all sufficiently large d we get

$$\max_{j=1, \dots, j_\delta} \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} < e^{-\tau b_d}) \leq \delta. \quad (27)$$

Also for all d such that $e^{-\tau b_d} < 1 - \delta$ we have

$$\begin{aligned} \max_{j=j_\delta, \dots, d} \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} < e^{-\tau b_d}) &\leq \max_{j=j_\delta, \dots, d} \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} / \bar{\lambda}_1^{X_{1,j}} < e^{-\tau b_d} / (1 - \delta)) \\ &\leq \max_{j=j_\delta, \dots, d} \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} / \bar{\lambda}_1^{X_{1,j}} < 1) \\ &\leq \max_{j=j_\delta, \dots, d} \sum_{k=2}^{\infty} \bar{\lambda}_k^{X_{1,j}}. \end{aligned}$$

Hence the condition (25) follows from (26) and (27).

Our next theorem gives the description of all possible functions q , which may appear in (6) under the assumption (25). Necessary notions and facts from probability theory can be found in Appendix.

Theorem 2 Let $(a_d)_{d \in \mathbb{N}}$ be a sequence, $(b_d)_{d \in \mathbb{N}}$ be a positive sequence such that $b_d \rightarrow \infty$, $d \rightarrow \infty$. Let a non-increasing function $q: (0, 1) \rightarrow \mathbb{R}$ and a distribution function G satisfy the equation $q(\varepsilon) = G^{-1}(1 - \varepsilon^2)$ for all $\varepsilon \in \mathbf{C}(q)$. Suppose that under the condition (25), the following asymptotics holds

$$\ln n^{X_d}(\varepsilon) = a_d + q(\varepsilon)b_d + o(b_d), \quad d \rightarrow \infty, \quad \text{for all } \varepsilon \in \mathbf{C}(q). \quad (28)$$

Then G is self-decomposable with zero Lévy spectral function on $(-\infty, 0)$. If G is non-degenerate, then $b_{d+1}/b_d \rightarrow 1$ as $d \rightarrow \infty$.

For proving this and next theorems we will introduce auxiliary random variables U_j , $j \in \mathbb{N}$, with the following distributions

$$\mathbb{P}(U_j = \lfloor \ln \bar{\lambda}_k^{X_{1,j}} \rfloor) = \mathbf{m}^{X_{1,j}}(\bar{\lambda}_k^{X_{1,j}}) \cdot \bar{\lambda}_k^{X_{1,j}}, \quad (29)$$

where $j \in \mathbb{N}$ and $k \in \mathbb{N}$ such that $\lambda_k^{X_{1,j}} > 0$.

Lemma 1 Let U_j , $j \in \mathbb{N}$, be independent random variables with distributions (29). Then we have

$$\sum_{k=1}^{\infty} \bar{\lambda}_k^{X_d} \mathbb{1}(\bar{\lambda}_k^{X_d} \geq e^{-x}) = \mathbb{P}\left(\sum_{j=1}^d U_j \leq x\right) \quad \text{for all } d \in \mathbb{N}, \quad x \in \mathbb{R}. \quad (30)$$

Proof of Lemma 1. Let us fix $d \in \mathbb{N}$ and $x \in \mathbb{R}$. If $x < 0$ then, by $\bar{\lambda}_k^{X_d} \leq 1$ and non-negativity of all U_j , $j \in \mathbb{N}$, both sides of (30) equal zero and hence it holds. Let $x \geq 0$. According to the multiplicative structure of $\bar{\lambda}_k^{X_d}$, $k \in \mathbb{N}$, we can write

$$\begin{aligned} \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_d} \mathbb{1}(\bar{\lambda}_k^{X_d} \geq e^{-x}) &= \sum_{k \in \mathbb{N}^d} \prod_{j=1}^d \bar{\lambda}_{k_j}^{X_{1,j}} \mathbb{1}\left(\prod_{j=1}^d \bar{\lambda}_{k_j}^{X_{1,j}} \geq e^{-x}\right) \\ &= \sum_{k \in \mathbb{N}^d} \prod_{j=1}^d \bar{\lambda}_{k_j}^{X_{1,j}} \mathbb{1}\left(\sum_{j=1}^d |\ln \bar{\lambda}_{k_j}^{X_{1,j}}| \leq x\right) \\ &= \sum_{\substack{j=1, \dots, d \\ \mu_j \in \mathcal{V}^{X_{1,j}}}} \prod_{j=1}^d \mathbf{m}^{X_{1,j}}(\mu_j) \mu_j \mathbb{1}\left(\sum_{j=1}^d |\ln \mu_j| \leq x\right), \end{aligned}$$

where $\mathcal{V}^{X_{1,j}} := \{\bar{\lambda}_k^{X_{1,j}} : k \in \mathbb{N}\}$ is the range of the sequence $(\bar{\lambda}_k^{X_{1,j}})_{k \in \mathbb{N}}$ for every $j \in \mathbb{N}$. By (29), we have

$$\begin{aligned} \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_d} \mathbb{1}(\bar{\lambda}_k^{X_d} \geq e^{-x}) &= \sum_{\substack{j=1, \dots, d \\ \mu_j \in \mathcal{V}^{X_{1,j}}}} \prod_{j=1}^d \mathbb{P}(U_j = |\ln \mu_j|) \mathbb{1}\left(\sum_{j=1}^d |\ln \mu_j| \leq x\right) \\ &= \sum_{\substack{j=1, \dots, d \\ \mu_j \in \mathcal{V}^{X_{1,j}}}} \mathbb{P}(U_j = |\ln \mu_j|, j \in \{1, \dots, d\}) \mathbb{1}\left(\sum_{j=1}^d |\ln \mu_j| \leq x\right) \\ &= \mathbb{P}\left(\sum_{j=1}^d U_j \leq x\right). \quad \square \end{aligned}$$

The described probabilistic construction was proposed by M. A. Lifshits and E. V. Tulyakova in the paper [18] in the context of approximation of tensor degrees of random elements, which have covariance operators with eigenvalues of unit multiplicity. We extend this approach to general tensor product-type random elements without any assumptions on marginal eigenvalue multiplicities.

Proof of Theorem 2. According to Theorem 1, the condition (28) is equivalent to the convergence

$$\lim_{d \rightarrow \infty} G_{a_d, b_d}^{X_d}(x) = G(x) \quad \text{for all } x \in \mathbf{C}(G), \quad (31)$$

where the functions $G_{a_d, b_d}^{X_d}$, $d \in \mathbb{N}$, are defined by (17) for given a_d , b_d , and X_d , $d \in \mathbb{N}$. Using Lemma 1, we can write

$$G_{a_d, b_d}^{X_d}(x) = \mathbb{P}\left(\frac{\sum_{j=1}^d U_j - a_d}{b_d} \leq x\right), \quad d \in \mathbb{N}, \quad x \in \mathbb{R}, \quad (32)$$

where the independent random variables U_j , $j \in \mathbb{N}$, are distributed according to (29) with the tails

$$\mathbb{P}(U_j > x) = \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} < e^{-x}), \quad x \geq 0.$$

From (25) and non-negativity of U_j , $j \in \mathbb{N}$, we have

$$\lim_{d \rightarrow \infty} \max_{j=1, \dots, d} \mathbb{P}(|U_j| > \tau b_d) = 0 \quad \text{for all } \tau > 0. \quad (33)$$

By Theorem 10 (see Appendix), the weak limit of $(G_{a_d, b_d}^{X_d})_{d \in \mathbb{N}}$, the function G , is necessarily self-decomposable. Let L denote the Lévy spectral function of G . From (A₂) of Theorem 11 and from non-negativity of U_j , $j \in \mathbb{N}$, we get $L(x) = 0$, $x < 0$. The assertion about $(b_d)_{d \in \mathbb{N}}$ follows directly from Theorem 10. \square

The next theorem provides a criterion for the asymptotics (28), where q is a quantile of self-decomposable law.

Theorem 3 *Let $(a_d)_{d \in \mathbb{N}}$ be a sequence, $(b_d)_{d \in \mathbb{N}}$ be a positive sequence such that $b_d \rightarrow \infty$, $d \rightarrow \infty$. Let a distribution function G is self-decomposable with triplet (γ, σ^2, L) , where $L(x) = 0$, $x < 0$. Let a non-increasing function $q: (0, 1) \rightarrow \mathbb{R}$ satisfy the equation $q(\varepsilon) = G^{-1}(1 - \varepsilon^2)$ for all $\varepsilon \in (0, 1)$. Under the condition (25), for the asymptotics*

$$\ln n^{X_d}(\varepsilon) = a_d + q(\varepsilon)b_d + o(b_d), \quad d \rightarrow \infty, \quad \text{for all } \varepsilon \in (0, 1), \quad (34)$$

the following ensemble of conditions is necessary and sufficient:

$$\begin{aligned} \text{(A)} \quad & \lim_{d \rightarrow \infty} \sum_{j=1}^d \sum_{k=1}^N \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} < e^{-xb_d}) = -L(x) \quad \text{for all } x > 0; \\ \text{(B)} \quad & \lim_{d \rightarrow \infty} \frac{1}{b_d} \left(\sum_{j=1}^d M_{1,N}^{X_{1,j}}(\tau b_d) - a_d \right) = \gamma + \int_0^\tau \frac{x^3 dL(x)}{1+x^2} - \int_\tau^\infty \frac{x dL(x)}{1+x^2} \quad \text{for all } \tau > 0; \\ \text{(C)} \quad & \lim_{\tau \rightarrow 0} \lim_{d \rightarrow \infty} \frac{1}{b_d^2} \sum_{j=1}^d \left(M_{2,N}^{X_{1,j}}(\tau b_d) - (M_{1,N}^{X_{1,j}}(\tau b_d))^2 \right) = \lim_{\tau \rightarrow 0} \lim_{d \rightarrow \infty} \dots = \sigma^2, \end{aligned}$$

where

$$M_{p,N}^{X_{1,j}}(x) := \sum_{k=1}^N |\ln \bar{\lambda}_k^{X_{1,j}}|^p \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} \geq e^{-x}), \quad x \geq 0, \quad p \in \{1, 2\},$$

and N is ∞ or any natural number such that

$$\sum_{j=1}^{\infty} \sum_{k=N+1}^{\infty} \bar{\lambda}_k^{X_{1,j}} < \infty. \quad (35)$$

Proof of Theorem 3. We first show that $\mathbf{C}(q) = (0, 1)$. If G is degenerate, then q is constant on $(0, 1)$ and, consequently, continuous. For non-degenerate case the assertion follows from Remarks 2 and 8 (see Appendix). Thus, by Theorem 1, the condition (34) is equivalent to the convergence (31), where $G_{a_d, b_d}^{X_d}$, $d \in \mathbb{N}$, are defined by (17) for given a_d , b_d , and X_d , $d \in \mathbb{N}$. According to (32), $G_{a_d, b_d}^{X_d}$, $d \in \mathbb{N}$, are distribution functions of centered and normalized sums of non-negative independent random variables U_j , $j \in \mathbb{N}$, satisfying (33). Consequently, for the convergence (31) the conditions

(A₁), (B), and (C) of Theorem 11 (see Appendix) are necessary and sufficient (we set $Y_j := U_j$, $j \in \mathbb{N}$, and $A_d := a_d$, $B_d := b_d$, $d \in \mathbb{N}$ in Theorem 11). For the case $N = \infty$ it directly yields the conditions (A), (B), and (C) of the theorem. Indeed, it is easily seen that

$$\mathbb{E} \left[U_j^p \mathbb{1}(|U_j| \leq x) \right] = M_{p,\infty}^{X_{1,j}}(x), \quad x \geq 0, \quad p \in \{1, 2\}.$$

Hence for any $x \geq 0$

$$\begin{aligned} \mathbf{Var} \left[U_j \mathbb{1}(|U_j| \leq x) \right] &= \mathbb{E} [\dots]^2 - (\mathbb{E} [\dots])^2 \\ &= M_{2,\infty}^{X_{1,j}}(x) - (M_{1,\infty}^{X_{1,j}}(x))^2. \end{aligned}$$

Suppose that there exists a natural number $N = N'$ that satisfies (35). Now we show that the conditions (A), (B), and (C) for $N = \infty$ are respectively equivalent to the same ones for $N = N'$.

For arbitrary small $\delta > 0$ we can find $j_\delta \in \mathbb{N}$ that yields $\sum_{j > j_\delta} \sum_{k > N'} \bar{\lambda}_k^{X_{1,j}} \leq \delta$. Thus for any $x > 0$ we have

$$\begin{aligned} \sum_{j=1}^d \sum_{k > N'} \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} < e^{-xb_d}) &\leq \sum_{j \leq j_\delta} \sum_{k > N'} \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} < e^{-xb_d}) + \sum_{j > j_\delta} \sum_{k > N'} \bar{\lambda}_k^{X_{1,j}} \\ &\leq j_\delta \cdot \max_{j=1, \dots, j_\delta} \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} < e^{-xb_d}) + \delta. \end{aligned}$$

By (25), the last expression can be made arbitrary small by the choice of sufficiently small δ and large d . This proves the equivalence of (A) for $N = \infty$ and (A) for $N = N'$ under the condition (35) with $N = N'$.

To continue the proof, we need the following relation

$$\lim_{d \rightarrow \infty} \frac{1}{b_d} \max_{j=1, \dots, d} \sum_{k=1}^{\infty} |\ln \bar{\lambda}_k^{X_{1,j}}| \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} \geq e^{-\tau b_d}) = 0 \quad \text{for all } \tau > 0, \quad (36)$$

which follows from (25) and the inequalities

$$\begin{aligned} \frac{1}{b_d} \max_{j=1, \dots, d} \sum_{k=1}^{\infty} |\ln \bar{\lambda}_k^{X_{1,j}}| \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} \geq e^{-\tau b_d}) &\leq \frac{1}{b_d} \max_{j=1, \dots, d} \left\{ \sum_{k=1}^{\infty} |\ln \bar{\lambda}_k^{X_{1,j}}| \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} \geq e^{-hb_d}) \right. \\ &\quad \left. + \sum_{k=1}^{\infty} |\ln \bar{\lambda}_k^{X_{1,j}}| \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} \in [e^{-\tau b_d}, e^{-hb_d}]) \right\} \\ &\leq h + \tau \max_{j=1, \dots, d} \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} < e^{-hb_d}), \end{aligned}$$

where $h < \tau$ is arbitrary small positive number.

Next, for any $\tau > 0$ we estimate the following sums

$$\begin{aligned}
\frac{1}{b_d} \sum_{j=1}^d \sum_{k>N'} |\ln \bar{\lambda}_k^{X_{1,j}}| \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} \geq e^{-\tau b_d}) &\leq \frac{1}{b_d} \sum_{j \leq j_\delta} \sum_{k>N'} |\ln \bar{\lambda}_k^{X_{1,j}}| \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} \geq e^{-\tau b_d}) \\
&\quad + \frac{1}{b_d} \sum_{j>j_\delta} \sum_{k>N'} |\ln \bar{\lambda}_k^{X_{1,j}}| \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} \geq e^{-\tau b_d}) \\
&\leq \frac{j_\delta}{b_d} \max_{j=1, \dots, j_\delta} \sum_{k=1}^{\infty} |\ln \bar{\lambda}_k^{X_{1,j}}| \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} \geq e^{-\tau b_d}) \\
&\quad + \tau \sum_{j>j_\delta} \sum_{k>N'} \bar{\lambda}_k^{X_{1,j}} \\
&\leq \frac{j_\delta}{b_d} \max_{j=1, \dots, d} \sum_{k=1}^{\infty} |\ln \bar{\lambda}_k^{X_{1,j}}| \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} \geq e^{-\tau b_d}) + \tau \delta.
\end{aligned}$$

Here the last expression can be made arbitrary small by the choice of sufficiently small δ and large d in view of (36). Thus we have the equivalence of (B) for $N = \infty$ and (B) for $N = N'$ under the condition (35) with $N = N'$.

To obtain the equivalence (C) for $N = \infty$ and (C) for $N = N'$, it is sufficient to show that

$$\lim_{d \rightarrow \infty} \frac{1}{b_d^2} \sum_{j=1}^d \sum_{k>N'} |\ln \bar{\lambda}_k^{X_{1,j}}|^2 \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} \geq e^{-\tau b_d}) = 0 \quad \text{for all } \tau > 0.$$

This follows from the estimate

$$\frac{1}{b_d^2} \sum_{k>N'} |\ln \bar{\lambda}_k^{X_{1,j}}|^2 \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} \geq e^{-\tau b_d}) \leq \frac{\tau}{b_d} \sum_{k>N'} |\ln \bar{\lambda}_k^{X_{1,j}}| \bar{\lambda}_k^{X_{1,j}} \mathbb{1}(\bar{\lambda}_k^{X_{1,j}} \geq e^{-\tau b_d})$$

and the previous conclusions. \square

4.3 Applications to tensor products of Euler integrated processes

Let us consider Gaussian random process $E_r(t)$, $t \in [0, 1]$, with zero mean and the following correlation function

$$\mathcal{K}^{E_r}(t, s) := \int_{[0,1]^r} \min\{t, x_1\} \min\{x_1, x_2\} \dots \min\{x_r, s\} dx_1 \dots dx_r,$$

where $t, s \in [0, 1]$ and r is a non-negative integer. Such process is called the *Euler integrated process*. It is connected with the standard Wiener process $W(t)$, $t \in [0, 1]$, by the following integration scheme:

$$E_r(t) = (-1)^{a_1 + \dots + a_r} \int_{a_r}^t \int_{a_{r-1}}^{t_{r-1}} \dots \int_0^{t_2} \int_1^{t_1} W(s) ds dt_1 \dots dt_{r-1}, \quad t \in [0, 1],$$

where $a_{2k-1} := 1$, $a_{2k} := 0$, $k \in \mathbb{N}$. The Euler integrated process is well known object: related boundary value problems were considered in [3], small ball probabilities were investigated in the papers [7] and [25].

Let us consider the sequence of the Euler integrated processes $E_{r_j}(t)$, $t \in [0, 1]$, $j \in \mathbb{N}$, with correlation functions $\mathcal{K}^{E_{r_j}}$, $j \in \mathbb{N}$, respectively. By $(r_j)_{j \in \mathbb{N}}$ we will always mean a non-decreasing sequence of non-negative integers. Consider a sequence of zero-mean random fields $\mathbb{E}_d(t)$, $t \in [0, 1]^d$, $d \in \mathbb{N}$, with the following correlation functions

$$\mathcal{K}^{\mathbb{E}_d}(t, s) = \prod_{j=1}^d \mathcal{K}^{E_{r_j}}(t_j, s_j), \quad t, s \in [0, 1]^d, \quad d \in \mathbb{N}.$$

We consider every process $E_{r_j}(t)$, $t \in [0, 1]$, as a random element E_{r_j} of the space $L_2([0, 1])$. The covariance operator $K^{E_{r_j}}$ of E_{r_j} is the integration operator with kernel $\mathcal{K}^{E_{r_j}}$. The eigenpairs of $K^{E_{r_j}}$ are exactly known (see [3] and [7]):

$$\lambda_k^{E_{r_j}} = \frac{1}{(\pi(k - 1/2))^{2r_j+2}}, \quad \psi_k^{E_{r_j}}(t) = \sqrt{2} \sin(\pi(k - 1/2)t), \quad k \in \mathbb{N}, \quad t \in [0, 1].$$

It is easily seen that every field $\mathbb{E}_d(t)$, $t \in [0, 1]^d$, as a random element \mathbb{E}_d of $L_2([0, 1]^d)$, has a covariance operator of the form $K^{\mathbb{E}_d} = \otimes_{j=1}^d K^{E_{r_j}}$, with eigenpairs (22) (we set $X_d = \mathbb{E}_d$, $X_{1,j} = E_{r_j}$). Therefore, by definition from Section 2, \mathbb{E}_d is a tensor product of E_{r_j} , i.e. $\mathbb{E}_d = \otimes_{j=1}^d E_{r_j}$.

Let us consider the sequence APP of approximation problems for \mathbb{E}_d , $d \in \mathbb{N}$. The criteria of all types of tractability for APP were obtained in the paper [17]. We recall a part of this result here.

Theorem 4 *APP is weakly tractable iff*

$$\lim_{j \rightarrow \infty} r_j = \infty; \tag{37}$$

APP is quasi-polynomially tractable iff

$$\sup_{d \in \mathbb{N}} \frac{1}{\ln_+ d} \sum_{j=1}^d (r_j + 1) 3^{-2r_j-2} < \infty; \tag{38}$$

APP is polynomially tractable iff it is strongly polynomially tractable iff

$$\sum_{j=1}^{\infty} 3^{-2\tau(r_j+1)} < \infty \quad \text{for some } \tau \in (0, 1). \tag{39}$$

In fact, depending on the type of tractability, the previous theorem provides a majorant (see (2)–(5)) for the approximation complexity $n^{\mathbb{E}_d}(\varepsilon)$ for all $\varepsilon \in (0, 1)$ and $d \in \mathbb{N}$.

We will investigate $n^{\mathbb{E}_d}(\varepsilon)$ for arbitrarily fixed $\varepsilon \in (0, 1)$ and $d \rightarrow \infty$. In order to compare our results with tractability bounds, we restrict ourselves only to quasi-polynomially tractable sequences APP with the following behaviour of $(r_j)_{j \in \mathbb{N}}$:

$$3^{-2r_j-2} \sim \frac{\beta}{j(\ln j)^p}, \quad j \rightarrow \infty, \tag{40}$$

where $\beta > 0$ and $p \geq 1$.

Remark 4 *Under the assumption (40) for some $\beta > 0$ and $p \in \mathbb{R}$, APP is weakly tractable, but not strongly polynomially tractable. APP is quasi-polynomially tractable iff $p \geq 1$.*

Indeed, weak tractability immediately follows from (37). It is easy to check that (39) does not hold. Consider the behaviour of sums from (38) as $d \rightarrow \infty$

$$\begin{aligned}
\frac{1}{\ln_+ d} \sum_{j=1}^d (1 + r_j) 3^{-2r_j-2} &\sim \frac{\beta}{(2 \ln 3) \ln d} \sum_{j=1}^d \frac{(\ln j)^{1-p}}{j} \\
&\sim \frac{\beta}{(2 \ln 3) \ln d} \int_2^d \frac{(\ln t)^{1-p}}{t} dt \\
&\sim \frac{\beta}{(2 \ln 3) \ln d} \cdot \left(\frac{(\ln d)^{2-p}}{2-p} \mathbb{1}(p \neq 2) + (\ln \ln d) \mathbb{1}(p = 2) \right) \\
&= \frac{\beta}{2 \ln 3} \cdot \left(\frac{(\ln d)^{1-p}}{2-p} \mathbb{1}(p \neq 2) + \frac{\ln \ln d}{\ln d} \mathbb{1}(p = 2) \right).
\end{aligned}$$

From this we conclude the second assertion of the remark.

Proposition 6 *Suppose that (40) holds for some $\beta > 0$ and $p \geq 1$. If $p > 1$ then*

$$\sup_{d \in \mathbb{N}} n^{\mathbb{E}_d}(\varepsilon) < \infty \quad \text{for all } \varepsilon \in (0, 1).$$

If $p = 1$ then

$$\lim_{d \rightarrow \infty} n^{\mathbb{E}_d}(\varepsilon) = \infty \quad \text{for all } \varepsilon \in (0, 1).$$

Proof of Proposition 6. Consider the following sum

$$\sum_{j=1}^d \sum_{k=2}^{\infty} \frac{\lambda_k^{E_{r_j}}}{\lambda_1^{E_{r_j}}} = \sum_{j=1}^d \sum_{k=2}^{\infty} \frac{1}{(2k-1)^{2r_j+2}} = \sum_{j=1}^d c(r_j) 3^{-2r_j-2},$$

where $c(r_j) := 1 + \sum_{k=3}^{\infty} ((2k-1)/3)^{-2r_j-2}$. Under the assumption (40) for some $\beta > 0$ and $p \geq 1$, $c(r_j) \rightarrow 1$, $j \rightarrow \infty$, and the series $\sum_{j=1}^{\infty} 3^{-2r_j-2}$ converges for $p > 1$ and it diverges for $p = 1$. Applying Proposition 5, we obtain the required assertions. \square

From this proposition we can see that for the cases with $p > 1$ the bound (3) for $n^{X_d}(\varepsilon)$ can be rather crude under the setting “ ε is fixed, $d \rightarrow \infty$ ”, because it depends on d , whereas $n^{X_d}(\varepsilon)$ is bounded on d .

For the case $p = 1$ we can find logarithmic asymptotics of the approximation complexity. Here the *convolution powers of the Dickman distribution* appear (see Appendix).

Proposition 7 *Under the assumption (40) with $p = 1$ and $\beta > 0$, we have*

$$\ln n^{\mathbb{E}_d}(\varepsilon) = D_{\beta}^{-1}(1 - \varepsilon^2) \ln d + o(\ln d), \quad d \rightarrow \infty, \quad \text{for all } \varepsilon \in (0, 1), \quad (41)$$

where D_{β} is the distribution function of β -convolution power of the Dickman distribution.

Proof of Proposition 7. Let us write the expression for traces of $K^{E_{r_j}}$ in the following form:

$$\Lambda^{E_{r_j}} = \sum_{k=1}^{\infty} \lambda_k^{E_{r_j}} = \sum_{k=1}^{\infty} (\pi(k-1/2))^{-2r_j-2} = \frac{1}{\omega_{r_j}} \cdot \left(\frac{2}{\pi}\right)^{2r_j+2},$$

where $\omega_{r_j} := (\sum_{k=1}^{\infty} (2k-1)^{-2r_j-2})^{-1}$, $j \in \mathbb{N}$. According to the notation, we have

$$\bar{\lambda}_k^{E_{r_j}} = \frac{\omega_{r_j}}{(2k-1)^{2r_j+2}}, \quad k \in \mathbb{N}, \quad j \in \mathbb{N}.$$

We see that $\bar{\lambda}_1^{E_{r_j}} = \omega_{r_j} \rightarrow 1$ as $j \rightarrow \infty$. Hence, by Remark 3, the condition (25) holds for any sequence $(b_d)_{d \in \mathbb{N}}$ such that $b_d \rightarrow \infty$, $d \rightarrow \infty$. Next, consider the sums

$$\sum_{j=1}^d \sum_{k=3}^{\infty} \bar{\lambda}_k^{E_{r_j}} = \sum_{j=1}^d \sum_{k=3}^{\infty} \frac{\omega_{r_j}}{(2k-1)^{2r_j+2}} \leq C \cdot \sum_{j=1}^d 5^{-2r_j-2},$$

where we set $C_1 := 1 + \sum_{k=4}^{\infty} (5/(2k-1))^2 < \infty$. Under our assumptions on $(r_j)_{j \in \mathbb{N}}$, we get

$$5^{-2r_j-2} \sim \left(\frac{\beta}{j \ln j}\right)^{\frac{\ln 5}{\ln 3}}, \quad j \rightarrow \infty,$$

which gives the convergence of the series $\sum_{j=1}^{\infty} \sum_{k=3}^{\infty} \bar{\lambda}_k^{E_{r_j}}$. Thus, in order to obtain the asymptotics (41), it is sufficient to check conditions (A)–(C) of Theorem 3 with $X_d = \mathbb{E}_d$, $X_{1,j} = E_{r_j}$, $G = D_\beta$, $a_d = 0$, $b_d = \ln_+ d$, and $N = 2$. Here D_β has the following triplet $\gamma := \beta\pi/4$, $\sigma^2 := 0$, $L(x) := \beta \ln x \mathbf{1}(x \in (0, 1])$ (see Appendix).

Let us check the condition (A) of Theorem 3:

$$\lim_{d \rightarrow \infty} \sum_{j=1}^d \sum_{k=1}^2 \bar{\lambda}_k^{E_{r_j}} \mathbf{1}(\bar{\lambda}_k^{E_{r_j}} < e^{-x \ln d}) = -\beta \ln x \mathbf{1}(x \in (0, 1]) \quad \text{for all } x > 0. \quad (42)$$

Since $\bar{\lambda}_1^{E_{r_j}} \rightarrow 1$ as $j \rightarrow \infty$, for any $x > 0$ and sufficiently large d we have $\min_{j \in \mathbb{N}} \bar{\lambda}_1^{E_{r_j}} > e^{-x \ln d}$ and, consequently,

$$\sum_{j=1}^d \sum_{k=1}^2 \bar{\lambda}_k^{E_{r_j}} \mathbf{1}(\bar{\lambda}_k^{E_{r_j}} < e^{-x \ln d}) = \sum_{j=1}^d \bar{\lambda}_2^{E_{r_j}} \mathbf{1}(\bar{\lambda}_2^{E_{r_j}} < e^{-x \ln d}).$$

The right-hand side of this equation tends to $-\beta \ln x \mathbf{1}(x \in (0, 1])$ as $d \rightarrow \infty$ for any $x > 0$, iff the sums $\sum_{j=1}^d 3^{-2r_j-2} \mathbf{1}(3^{2r_j+2} > e^{x \ln d})$ tend to the same limit as $d \rightarrow \infty$ for any $x > 0$. This follows from the equality $\bar{\lambda}_2^{E_{r_j}} = \omega_{r_j} 3^{-2r_j-2}$ with $\omega_{r_j} \rightarrow 1$, $j \rightarrow \infty$, and from the continuity of the limit function. Thus (42) is equivalent to

$$\lim_{d \rightarrow \infty} \sum_{j=1}^d 3^{-2r_j-2} \mathbf{1}(3^{2r_j+2} > e^{x \ln d}) = -\beta \ln x \mathbf{1}(x \in (0, 1]) \quad \text{for all } x > 0. \quad (43)$$

For $x \geq 1$ we have $3^{2r_j+2} > e^{x \ln d}$, $j = 1, \dots, d$, for all sufficiently large d . Therefore (43) obviously holds in this case. For $x \in (0, 1)$ we set $j_{d,x} := \min\{j \in \mathbb{N} : 3^{2r_j+2} > e^{x \ln d}\}$. Using Lemmas 5 and 6 from Appendix, we have

$$j_{d,x} \sim \beta d^x (\ln)^\#(\beta d^x) \sim \frac{\beta d^x}{x \ln d}, \quad d \rightarrow \infty. \quad (44)$$

Under our assumptions on $(r_j)_{j \in \mathbb{N}}$, we have

$$\sum_{j=1}^d 3^{-2r_j-2} \mathbb{1}(3^{2r_j+2} > e^{x \ln d}) = \sum_{j=j_{d,x}}^d 3^{-2r_j-2} \sim \sum_{j=j_{d,x}}^d \frac{\beta}{j \ln j}, \quad d \rightarrow \infty.$$

Using (44) and the asymptotics (see [23], 2.13, p. 21)

$$\sum_{k=1}^n \frac{1}{k \ln k} = \ln \ln n + c + o(1), \quad n \rightarrow \infty, \quad (45)$$

with some constant c , we obtain

$$\begin{aligned} \sum_{j=j_{d,x}}^d \frac{\beta}{j \ln j} &= \beta (\ln \ln d - \ln \ln j_{d,x}) + o(1) \\ &= \beta (\ln \ln d - \ln \ln d^x) + o(1) \\ &= -\beta \ln x + o(1), \end{aligned}$$

as $d \rightarrow \infty$. Thus we have the convergence (43).

Next, we check the condition (B) of Theorem 3:

$$\lim_{d \rightarrow \infty} \frac{1}{\ln d} \sum_{j=1}^d \sum_{k=1}^2 |\ln \bar{\lambda}_k^{E_{r_j}}| \bar{\lambda}_k^{E_{r_j}} \mathbb{1}(\bar{\lambda}_k^{E_{r_j}} \geq e^{-\tau \ln d}) = \frac{\beta \pi}{4} + \int_0^{\min\{\tau, 1\}} \frac{\beta x^2 dx}{1+x^2} - \int_{\min\{\tau, 1\}}^1 \frac{\beta dx}{1+x^2}$$

for all $\tau > 0$. Here the right-hand side exactly equals to $\beta \min\{\tau, 1\}$. In the left-hand side we estimate

$$\sum_{j=1}^d |\ln \bar{\lambda}_1^{E_{r_j}}| \bar{\lambda}_1^{E_{r_j}} \leq \sum_{j=1}^d (1 - \bar{\lambda}_1^{E_{r_j}}) = \sum_{j=1}^d \sum_{k=2}^{\infty} \frac{\omega_{r_j}}{(2k-1)^{2r_j+2}} \leq C_2 \sum_{j=1}^d 3^{-2r_j-2},$$

where we set $C_2 := 1 + \sum_{k=3}^{\infty} (3/(2k-1))^2 < \infty$. Next, by (45), we see that

$$\frac{C_2}{\ln d} \sum_{j=1}^d 3^{-2r_j-2} \sim \frac{C_2}{\ln d} \sum_{j=1}^d \frac{\beta}{j \ln j} \sim \frac{C_2 \beta \ln \ln d}{\ln d} \rightarrow 0, \quad d \rightarrow \infty.$$

Hence we only need to prove the convergence

$$\lim_{d \rightarrow \infty} \frac{1}{\ln d} \sum_{j=1}^d |\ln \bar{\lambda}_2^{E_{r_j}}| \bar{\lambda}_2^{E_{r_j}} \mathbb{1}(\bar{\lambda}_2^{E_{r_j}} \geq e^{-\tau \ln d}) = \beta \min\{\tau, 1\} \quad \text{for all } \tau > 0. \quad (46)$$

Since $\bar{\lambda}_2^{E_{r_j}} = \omega_{r_j} 3^{-2r_j-2}$ with $\omega_{r_j} \rightarrow 1$, $j \rightarrow \infty$, and the limit function $\tau \mapsto \beta \min\{\tau, 1\}$ is continuous at any $\tau > 0$, (46) is equivalent to

$$\lim_{d \rightarrow \infty} \frac{1}{\ln d} \sum_{\substack{j=1, \dots, d \\ j < j_{d,\tau}}} \ln(3^{2r_j+2}) 3^{-2r_j-2} = \beta \min\{\tau, 1\} \quad \text{for all } \tau > 0.$$

But this follows from (44) and the following equivalences

$$\sum_{\substack{j=1, \dots, d \\ j < j_{d,\tau}}} \ln(3^{2r_j+2}) 3^{-2r_j-2} \sim \sum_{\substack{j=1, \dots, d \\ j < j_{d,\tau}}} \frac{\beta}{j} \sim \beta \ln \min\{d, j_{d,\tau}\} \sim \beta \min\{\tau, 1\} \ln d, \quad d \rightarrow \infty.$$

Now we check the condition (C) of Theorem 3. It is sufficient to prove

$$\lim_{\tau \rightarrow 0} \lim_{d \rightarrow \infty} \sum_{j=1}^d \sum_{k=1}^2 \frac{|\ln \bar{\lambda}_k^{E_{r_j}}|^2}{(\ln d)^2} \bar{\lambda}_k^{E_{r_j}} \mathbf{1}(\bar{\lambda}_k^{E_{r_j}} \geq e^{-\tau \ln d}) = 0.$$

But this follows from (46) and the next bound

$$\begin{aligned} \lim_{d \rightarrow \infty} \sum_{j=1}^d \sum_{k=1}^2 \frac{|\ln \bar{\lambda}_k^{E_{r_j}}|^2}{(\ln d)^2} \bar{\lambda}_k^{E_{r_j}} \mathbf{1}(\bar{\lambda}_k^{E_{r_j}} \geq e^{-\tau \ln d}) &\leq \lim_{d \rightarrow \infty} \sum_{j=1}^d \sum_{k=1}^2 \frac{\tau |\ln \bar{\lambda}_k^{E_{r_j}}|}{\ln d} \bar{\lambda}_k^{E_{r_j}} \mathbf{1}(\bar{\lambda}_k^{E_{r_j}} \geq e^{-\tau \ln d}) \\ &= \beta \tau \min\{\tau, 1\}. \end{aligned}$$

Thus we have checked the conditions (A)–(C) of Theorem 3. Hence, by this theorem we have the required asymptotics. \square

5 Approximation of tensor degree-type random elements

In this section we consider sequences of tensor degree-type random elements. Here we deal with an important particular case of the linear tensor approximation problems, which were considered in Section 4. Let X be a zero-mean random element of a separable Hilbert space H and assume that $\mathbb{E} \|X\|_H^2 < \infty$. Let $X_d = X^{\otimes d}$ for every $d \in \mathbb{N}$. From (22) we see that eigenvalues and eigenvectors of K^{X_d} are respectively the following:

$$\prod_{j=1}^d \lambda_{k_j}^X, \quad \bigotimes_{j=1}^d \psi_{k_j}^X, \quad k_1, k_2, \dots, k_d \in \mathbb{N}.$$

By formula (23), for traces of K^{X_d} we have $\Lambda^{X_d} = (\Lambda^X)^d$, $d \in \mathbb{N}$. As in Section 4, we always assume that $\lambda_1^X > 0$.

Let us proceed with some elementary remarks following from the definitions. If $\lambda_1^X = \Lambda^X$ then by formula (9) we have $n^{X_d}(\varepsilon) = 1$ for all $\varepsilon \in (0, 1)$ and $d \in \mathbb{N}$. In more interesting case, when $\lambda_1^X < \Lambda^X$, from the equality $\lambda_1^{X_d} = (\lambda_1^X)^d$ and lower bound (11) we obtain the inequality

$$n^{X_d}(\varepsilon) \geq (1 - \varepsilon^2) (\Lambda^X / \lambda_1^X)^d \quad \text{for all } \varepsilon \in (0, 1), \quad d \in \mathbb{N}. \quad (47)$$

Therefore $n^{X_d}(\varepsilon) \rightarrow \infty$ as $d \rightarrow \infty$ at least exponentially for all fixed $\varepsilon \in (0, 1)$ i.e. here we always have the curse of dimensionality.

In general, bound (47) can not be improved. Indeed, consider the case when

$$\exists l^X \in \mathbb{N} \quad \lambda_k^X = \lambda_1^X \mathbf{1}(k \leq l^X), \quad k \in \mathbb{N}. \quad (48)$$

Under this condition, we have $l^X = \Lambda^X / \lambda_1^X$ and $\lambda_k^{X_d} = \lambda_1^{X_d} \mathbf{1}(k \leq (l^X)^d)$, $k \in \mathbb{N}$. By the formula (10), for all $\varepsilon \in (0, 1)$ and $d \in \mathbb{N}$ it is easy to obtain the following inequalities

$$\sum_{k=1}^{n^{X_d}(\varepsilon)-1} \lambda_k^{X_d} < (1 - \varepsilon^2) \Lambda^{X_d} \leq \sum_{k=1}^{n^{X_d}(\varepsilon)} \lambda_k^{X_d},$$

which are reduced to the required bounds

$$n^{X_d}(\varepsilon) - 1 < (1 - \varepsilon^2) (\Lambda^X / \lambda_1^X)^d \leq n^{X_d}(\varepsilon) \quad \text{for all } \varepsilon \in (0, 1), d \in \mathbb{N},$$

i.e. $n^{X_d}(\varepsilon) = \lceil (1 - \varepsilon^2) (\Lambda^X / \lambda_1^X)^d \rceil$, where $\lceil \cdot \rceil$ is a ceiling function.

5.1 Logarithmic asymptotics of the approximation complexity

For tensor degree-type random elements stable distributions (see Appendix) play a crucial role in approximation problems, as the following theorem shows.

Theorem 5 *Let $(a_d)_{d \in \mathbb{N}}$ be a sequence, $(b_d)_{d \in \mathbb{N}}$ be a positive sequence such that $b_d \rightarrow \infty$, $d \rightarrow \infty$. Let a non-increasing function $q: (0, 1) \rightarrow \mathbb{R}$ and a distribution function G satisfy the equation $q(\varepsilon) = G^{-1}(1 - \varepsilon^2)$ for all $\varepsilon \in \mathbf{C}(q)$. Suppose that we have*

$$\ln n^{X_d}(\varepsilon) = a_d + q(\varepsilon)b_d + o(b_d), \quad d \rightarrow \infty, \quad \text{for all } \varepsilon \in \mathbf{C}(q). \quad (49)$$

Then G is a stable distribution function. If G is non-degenerate, then $b_{d+1}/b_d \rightarrow \infty$, $d \rightarrow \infty$. If G is non-degenerate and non-normal, then $G = S_{\alpha, \rho, 1, \mu}$ for some $\alpha \in (0, 2)$, $\rho > 0$, $\mu \in \mathbb{R}$.

Proof of Theorem 5. By Theorem 1, the condition (49) is equivalent to the convergence (31), where the functions $G_{a_d, b_d}^{X_d}$, $d \in \mathbb{N}$, are defined by (17) for given a_d , b_d , and X_d , $d \in \mathbb{N}$. According to Lemma 1, $G_{a_d, b_d}^{X_d}(x)$ admits the representation (32), where the independent random variables U_j , $j \in \mathbb{N}$, have the common distribution:

$$\mathbb{P}(U_j = \lfloor \ln \bar{\lambda}_k^X \rfloor) = \mathbf{m}^X(\bar{\lambda}_k^X) \cdot \bar{\lambda}_k^X, \quad j \in \mathbb{N}, \quad (50)$$

where $k \in \mathbb{N}$ is such that $\lambda_k^X > 0$. By Theorem 12 (see Appendix), the weak limit of $G_{a_d, b_d}^{X_d}(x)$ is necessarily stable.

It is easy to check that the condition (25) is satisfied for tensor degree-type random elements (set $X_{1,j} := X$) for any positive sequence $(b_d)_{d \in \mathbb{N}}$ such that $b_d \rightarrow \infty$, $d \rightarrow \infty$. Then Theorem 2 gives the required assertion for $(b_d)_{d \in \mathbb{N}}$.

Assume that $G = S_{\alpha, \rho, \beta, \mu}$ for some $\alpha \in (0, 2)$, $\rho > 0$, $\beta \in [-1, 1]$, and $\mu \in \mathbb{R}$. Let L denote its Lévy spectral function. From Theorem 2 we have $L(x) = 0$ for all $x < 0$. In view of formula (72), we conclude that $\beta = 1$. \square

According to the previous theorem, we can consider only functions q such that $q(\varepsilon) = G^{-1}(1 - \varepsilon^2)$ for all $\varepsilon \in \mathbf{C}(q)$, where G is a stable distribution function. Without loss of generality, we will restrict ourselves to the cases, where G is non-degenerate and it also has the standard form (see (71) and (72)).

Before formulating the criteria for (49), we consider conditions that appear there. The wide class of cases corresponds to the assumption:

$$\sum_{k \in \mathbb{N}: \bar{\lambda}_k^X > 0} |\ln \bar{\lambda}_k^X|^2 \bar{\lambda}_k^X < \infty. \quad (51)$$

We also consider marginal random elements X with the following regular variation of $\bar{\lambda}_k^X$, $k \in \mathbb{N}$:

$$\sum_{k=1}^{\infty} \bar{\lambda}_k^X \mathbf{1}(\bar{\lambda}_k^X < e^{-x}) = x^{-\alpha} \varphi(x), \quad (52)$$

where $\alpha \geq 0$ and φ is some slowly varying function at infinity (SVF for short, see Appendix). Under the assumption (52) with some $\alpha > 2$, the condition (51) always holds. If either (51) or (52) with $\alpha > 1$ hold, then the entropy of $\bar{\lambda}_k^X$, $k \in \mathbb{N}$, is well defined:

$$E^X := \sum_{k \in \mathbb{N}: \bar{\lambda}_k^X > 0} |\ln \bar{\lambda}_k^X| \bar{\lambda}_k^X. \quad (53)$$

Under the assumption (51), the following deviation characteristic is also important:

$$\sigma^X := \left(\sum_{k \in \mathbb{N}: \bar{\lambda}_k^X > 0} \left(|\ln \bar{\lambda}_k^X| - E^X \right)^2 \bar{\lambda}_k^X \right)^{1/2}. \quad (54)$$

The next theorem provides a criterion of the asymptotics (49), where q is a quantile of the distribution function Φ of the standard normal law.

Theorem 6 *Let $(a_d)_{d \in \mathbb{N}}$ be a sequence, $(b_d)_{d \in \mathbb{N}}$ be a positive sequence such that $b_d \rightarrow \infty$, $d \rightarrow \infty$. For the asymptotics*

$$\ln n^{X_d}(\varepsilon) = a_d + \Phi^{-1}(1 - \varepsilon^2)b_d + o(b_d), \quad d \rightarrow \infty, \quad \text{for all } \varepsilon \in (0, 1), \quad (55)$$

it is necessary and sufficient to have:

- (i) *the condition (51) with $\sigma^X > 0$ or the condition (52) with $\alpha = 2$ and some SVF φ ;*
- (ii) *$a_d = E^X d + o(b_d)$, $d \rightarrow \infty$;*
- (iii) *$\lim_{d \rightarrow \infty} \frac{d}{b_d^2} (M_2^X(b_d) - M_1^X(b_d)^2) = 1$,*

where

$$M_p^X(x) := \sum_{k=1}^{\infty} |\ln \bar{\lambda}_k^X|^p \bar{\lambda}_k^X \mathbf{1}(\bar{\lambda}_k^X \geq e^{-x}), \quad x \geq 0, \quad p \in \{1, 2\}.$$

If (51) holds with $\sigma^X > 0$ then (iii) is equivalent to $b_d \sim \sigma^X d^{1/2}$, $d \rightarrow \infty$. If (52) holds with $\alpha = 2$ and SVF φ but (51) fails, then (iii) is equivalent to $b_d \sim d^{1/2} \varphi_2(d)$, $d \rightarrow \infty$, with SVF φ_2 , which is defined by $\varphi_2(d) := \sqrt{2} (1/\sqrt{\hat{\varphi}})^\#(d^{1/2})$, $d \in \mathbb{N}$, where $(\cdot)^\#$ is the de Bruijn conjugation² and SVF $\hat{\varphi}$ is defined by

$$\hat{\varphi}(x) := \int_0^x \frac{\varphi(t)}{t} dt, \quad x \geq 0. \quad (56)$$

Proof of Theorem 6. The function Φ is absolutely continuous and strictly increasing on \mathbb{R} . Hence, by Theorem 1, the condition (55) is equivalent to the convergence

$$\lim_{d \rightarrow \infty} G_{a_d, b_d}^{X_d}(x) = \Phi(x) \quad \text{for all } x \in \mathbb{R}, \quad (57)$$

where the functions $G_{a_d, b_d}^{X_d}$, $d \in \mathbb{N}$, are defined by (17) for given a_d , b_d , and X_d , $d \in \mathbb{N}$. By (32), $G_{a_d, b_d}^{X_d}(x)$, $d \in \mathbb{N}$, are distribution functions of centered and normalized sums of independent random variables U_j , $j \in \mathbb{N}$, with the common distribution (50).

Let us show the sufficiency of the assumptions (i)–(iii) for the convergence (57). In probabilistic interpretation the assumption (51) means $\mathbb{E} U_j^2 < \infty$. Also we can rewrite the condition (52) for $\alpha = 2$ as follows: $\mathbb{P}(U_j > x) = x^{-2} \varphi(x)$. Since $U_j \geq 0$, we have $\mathbb{P}(U_j < -x) = 0$ for all $x > 0$. Then the conditions (i) and (iii) are sufficient for the convergence $\lim_{d \rightarrow \infty} G_{E^X d, b_d}^{X_d}(x) = \Phi(x)$, $x \in \mathbb{R}$, by Theorem 13. Therefore, using the assumption (ii) of the theorem and the assertion (ii) of Lemma 3 we obtain (57).

Let us show the necessity of (i)–(iii). Under the convergence (57), the condition (i) holds by Theorem 13. Also we have $\lim_{d \rightarrow \infty} G_{E^X d, b_d^*}^{X_d}(x) = \Phi(x)$ for any $x \in \mathbb{R}$ and some sequence $(b_d^*)_{d \in \mathbb{N}}$ that satisfies (iii). According to the assertion (i) of Lemma 3, we obtain $b_d^* \sim b_d$ and $a_d = E^X d + o(b_d^*)$, $d \rightarrow \infty$, that directly yields (ii). The condition (iii) for $(b_d)_{d \in \mathbb{N}}$ follows from the slow variation of the function $x \mapsto M_2^X(x) - M_1^X(x)^2$, $x \geq 0$ (which is justified the next notes).

Under the condition (51), we have $M_2^X(x) - M_1^X(x)^2 \rightarrow \sigma^X$, $x \rightarrow \infty$, i.e. (iii) is equivalent to $b_d \sim \sigma^X d^{1/2}$, $d \rightarrow \infty$. In the remainder of this proof we assume (52) with $\alpha = 2$ and some SVF φ but (51) fails. Here we have $M_2^X(x) \rightarrow \infty$, $x \rightarrow \infty$, and, by the remark (2.6.14) from [13] (see p. 80), $M_1^X(x)^2 = o(M_2^X(x))$, $x \rightarrow \infty$. Hence (iii) is equivalent to $b_d^2 \sim d M_2^X(b_d)$, $d \rightarrow \infty$. Using integral representation for $M_2^X(x)$ and integrating it by parts, we get

$$M_2^X(x) = \int_0^x t^2 d \left(\sum_{k=1}^{\infty} \bar{\lambda}_k^X \mathbf{1}(\bar{\lambda}_k^X \geq e^{-t}) \right) = \int_0^x t^2 d(1 - t^{-2} \varphi(t)) = -\varphi(x) + 2\hat{\varphi}(x), \quad x \geq 0,$$

where $\hat{\varphi}$ is defined by (56). By Lemma 7 (see Appendix), $\hat{\varphi}$ is a SVF and also $\varphi(x) = o(\hat{\varphi}(x))$ as $x \rightarrow \infty$. Hence (iii) is equivalent to $b_d^2 \sim 2d \hat{\varphi}(b_d)$, $d \rightarrow \infty$. Rewriting this in the equivalent form

$$d \sim (b_d/\sqrt{2})^2 \cdot (1/\sqrt{\hat{\varphi}})^2(b_d/\sqrt{2}), \quad d \rightarrow \infty,$$

and using Lemma 5 from Appendix, we obtain the required relation for b_d , $d \in \mathbb{N}$. \square

²see Appendix.

Using tools of regular variation theory (see [2]), it is possible to find simpler asymptotic versions of the function φ_2 from Theorem 6 under special assumptions (see Lemma 6 and examples in the next subsection).

The previous theorem has the following important corollary concerning wide class of tensor degree-type random elements.

Theorem 7 *Under the assumption (51) with $\sigma^X > 0$, we have*

$$\ln n^{X_d}(\varepsilon) = E^X d + \Phi^{-1}(1 - \varepsilon^2)\sigma^X d^{1/2} + o(d^{1/2}), \quad d \rightarrow \infty, \quad \text{for all } \varepsilon \in (0, 1).$$

In fact, this theorem was obtained by M. A. Lifshits and E. V. Tulyakova in the paper [18]. However, strictly speaking, the proof from [18] was done only for sequences $(\bar{\lambda}_k^X)_{k \in \mathbb{N}}$ with unit multiplicity of every element, i.e. $\mathbf{m}^X(\bar{\lambda}_k^X) = 1$, $k \in \mathbb{N}$ (it is hidden in the last formula on p. 106 in [18]).

Next remarks follow directly from the definitions of E^X and σ^X .

Remark 5 *The condition (48) holds iff $\sigma^X = 0$.*

Remark 6 *The equality $\lambda_1^X = \Lambda^X$ (i.e. (48) with $l^X = 1$) holds iff $E^X = 0$.*

These remarks show that there is no loss of generality for us in assuming $\sigma^X > 0$ in Theorems 6 and 7. Also we can now conclude that under the assumption (i) of the previous theorem, the complexity $n^{X_d}(\varepsilon)$ grows mainly *exponentially* with the constant $E^X > 0$ as $d \rightarrow \infty$.

The next theorem provides a criterion of the asymptotics (49) with $G = S_{\alpha,1}$, $\alpha \in (0, 2)$.

Theorem 8 *Let $(a_d)_{d \in \mathbb{N}}$ be a sequence, $(b_d)_{d \in \mathbb{N}}$ be a positive sequence such that $b_d \rightarrow \infty$, $d \rightarrow \infty$. For the asymptotics*

$$\ln n^{X_d}(\varepsilon) = a_d + S_{\alpha,1}^{-1}(1 - \varepsilon^2)b_d + o(b_d), \quad d \rightarrow \infty, \quad \text{for all } \varepsilon \in (0, 1). \quad (58)$$

It is necessary and sufficient to have:

- (i) *the condition (52) with given α and some SVF φ ;*
- (ii) *$a_d = a_d^* + o(b_d)$, $d \rightarrow \infty$;*
- (iii) *$\lim_{d \rightarrow \infty} d \sum_{k=1}^{\infty} \bar{\lambda}_k^X \mathbb{1}(\bar{\lambda}_k^X < e^{-b_d}) = 1$.*

Here $a_d^ := 0$ for $\alpha \in (0, 1)$, $a_d^* := dE^X$ for $\alpha \in (1, 2)$ and*

$$a_d^* := d \sum_{k=1}^{\infty} |\ln \bar{\lambda}_k^X| \bar{\lambda}_k^X \mathbb{1}(\bar{\lambda}_k^X > e^{-b_d}) + (1 - \mathcal{C})b_d, \quad \text{for } \alpha = 1,$$

where \mathcal{C} is the Euler constant. Under the assumption (i) with some $\alpha \in (0, 2)$, the condition (iii) is equivalent to $b_d \sim d^{1/\alpha} \varphi_\alpha(d)$, $d \rightarrow \infty$, with SVF φ_α , defined by $\varphi_\alpha(d) := ((1/\varphi)^{1/\alpha})^\#(d^{1/\alpha})$, $d \in \mathbb{N}$, where $(\cdot)^\#$ is the de Bruijn conjugation.

Proof of Theorem 8. Since $S_{\alpha,1}$ is self-decomposable (see Appendix), the distribution function $S_{\alpha,1}$ is absolutely continuous on \mathbb{R} and strictly increasing on $(\text{lex} S_{\alpha,1}, \text{rex} S_{\alpha,1})$ in view of Remark 8. Hence, by Theorem 1, the condition (58) is equivalent to the convergence

$$\lim_{d \rightarrow \infty} G_{a_d, b_d}^{X_d}(x) = S_{\alpha,1}(x) \quad \text{for all } x \in \mathbb{R}, \quad (59)$$

where the functions $G_{a_d, b_d}^{X_d}$, $d \in \mathbb{N}$, are defined by (17) for given a_d , b_d , and X_d , $d \in \mathbb{N}$. By probabilistic representation (32), $G_{a_d, b_d}^{X_d}(x)$, $d \in \mathbb{N}$, are distribution functions of centered and normalized sums of independent random variables U_j , $j \in \mathbb{N}$, with the common distribution (50).

Let us show the sufficiency of the assumptions (i)–(iii) for the convergence (59). In probabilistic interpretation the assumption (52) can be written as follows: $\mathbb{P}(U_j > x) = x^{-\alpha} \varphi(x)$, $x \geq 0$. Since $U_j \geq 0$, we have $\mathbb{P}(U_j < -x) = 0$ for all $x > 0$. By Theorem 14, the conditions (i) and (iii) are sufficient for the convergence $\lim_{d \rightarrow \infty} G_{a_d^*, b_d^*}^{X_d}(x) = S_{\alpha,1}(x)$, $x \in \mathbb{R}$. Therefore, using the assumption (ii) of the theorem and the assertion (ii) of Lemma 3 we obtain (59).

Let us show the necessity of (i)–(iii). Under the convergence (59), the condition (i) holds by Theorem 14. Also we have $\lim_{d \rightarrow \infty} G_{a_d^*, b_d^*}^{X_d}(x) = S_{\alpha,1}(x)$ for any $x \in \mathbb{R}$ and some sequence $(b_d^*)_{d \in \mathbb{N}}$ that satisfies (iii). According to the assertion (i) of Lemma 3, we obtain $b_d^* \sim b_d$ and $a_d = a_d^* + o(b_d^*)$, $d \rightarrow \infty$, that yields (ii). The condition (iii) for $(b_d)_{d \in \mathbb{N}}$ follows from the regular variation of the function $x \mapsto \sum_{k=1}^{\infty} \lambda_k^X \mathbb{1}(\lambda_k^X < e^{-x})$, $x \geq 0$.

Under the condition (52) with some $\alpha \in (0, 2)$ and some SVF φ , (iii) can be rewritten as $d \sim (b_d)^\alpha (1/\varphi)^{1/\alpha}(b_d)^\alpha$, $d \rightarrow \infty$. Using Lemma 5 from Appendix, we obtain the required relation for b_d , $d \in \mathbb{N}$. \square

Simpler asymptotic versions of b_d can be obtained using Lemma 6 from Appendix.

Remark 7 Under the assumption (52) with $\alpha = 1$, (ii), and (iii), we have $b_d = o(a_d)$, $d \rightarrow \infty$.

Indeed, using integral representation for a_d and the condition (iii), we obtain as $d \rightarrow \infty$

$$\begin{aligned} a_d &= d \int_0^{b_d} t \, d\left(\sum_{k=1}^{\infty} \lambda_k^X \mathbb{1}(\lambda_k^X \geq e^{-t})\right) + (1 - \mathcal{C})b_d + o(b_d) \\ &= d \int_0^{b_d} t \, d(1 - t^{-1} \varphi(t)) + (1 - \mathcal{C})b_d + o(b_d) \\ &= -d\varphi(b_d) + d\hat{\varphi}(b_d) + (1 - \mathcal{C})b_d + o(b_d) \\ &= d\hat{\varphi}(b_d) - \mathcal{C}b_d + o(b_d) \\ &= b_d(\hat{\varphi}(b_d)/\varphi(b_d)) - \mathcal{C}b_d + o(b_d). \end{aligned}$$

By Lemma 7 (see Appendix), we have $\varphi(x) = o(\hat{\varphi}(x))$ as $x \rightarrow \infty$, which completes the verification.

Let us analyze the behaviour of $n^{X_d}(\varepsilon)$ as $d \rightarrow \infty$. Under the assumption (i) with $\alpha \in (1, 2)$, the approximation complexity grows *exponentially* due to a_d with *subexponential* term containing b_d . In the boundary case $\alpha = 1$ the main term a_d can yield more than exponential growth (see example in

the next subsection). Under the cases $\alpha \in (0, 1)$, factor $d^{1/\alpha}$ (of b_d) gives the *superexponential* growth of the quantity $n^{X_d}(\varepsilon)$.

Consider the unexplored pathological case when the assumption (52) is satisfied for $\alpha = 0$ and some SVF φ that $\varphi(x) \rightarrow 0$ as $x \rightarrow \infty$. Here it is impossible to find $(a_d)_{d \in \mathbb{N}}$ and $(b_d)_{d \in \mathbb{N}}$ for obtaining the asymptotics (49) with non-degenerate distribution function G (see comments in Subsection 6.3 in Appendix). Nevertheless, we obtain the following result.

Theorem 9 *If (52) holds with $\alpha = 0$ and some SVF φ , then*

$$\lim_{d \rightarrow \infty} d \varphi(\ln n^{X_d}(\varepsilon)) = -\ln(1 - \varepsilon^2) \quad \text{for all } \varepsilon \in (0, 1). \quad (60)$$

Proof of Theorem 9. At first we will prove that

$$\lim_{d \rightarrow \infty} \left[d \varphi\left(|\ln \bar{\lambda}^{X_d}(\varepsilon)|\right) \right] = -\ln(1 - \varepsilon^2) \quad \text{for all } \varepsilon \in (0, 1), \quad (61)$$

where $\bar{\lambda}^{X_d}(\varepsilon)$ is defined by (12). Let U_j , $j \in \mathbb{N}$, be independent random variables with common distribution (50). Then we have

$$\mathbb{P}(U_1 > x) = \sum_{k=1}^{\infty} \bar{\lambda}_k^X \mathbb{1}(\bar{\lambda}_k^X < e^{-x}) = \varphi(x), \quad x \geq 0.$$

Introduce the function $\tilde{\varphi}$:

$$\tilde{\varphi}(x) := 1 - \mathbb{E} \exp\{-U_1/x\} = 1 - \sum_{k=1}^{\infty} (\bar{\lambda}_k^X)^{1+1/x}, \quad x > 0.$$

By remarks to Theorem 15 (see Appendix), the function $\tilde{\varphi}$ is continuous and strictly decreasing on $(0, \infty)$. Also it satisfies:

$$\tilde{\varphi}(x) \sim \varphi(x), \quad x \rightarrow \infty. \quad (62)$$

Let us set

$$F^{X_d}(x) := \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_d} \mathbb{1}\left(-d \tilde{\varphi}(|\ln \bar{\lambda}_k^{X_d}|) \leq x\right), \quad x \in \mathbb{R}.$$

According to Lemma 1, F^{X_d} , $d \in \mathbb{N}$, have the following representations:

$$F^{X_d}(x) = \mathbb{P}\left(-d \tilde{\varphi}\left(\sum_{j=1}^d U_j\right) \leq x\right), \quad x \in \mathbb{R}.$$

By Theorem 15, for $x < 0$ we get:

$$F^{X_d}(x) = 1 - \mathbb{P}\left(d \tilde{\varphi}\left(\sum_{j=1}^d U_j\right) < -x\right) \rightarrow e^x, \quad d \rightarrow \infty.$$

From this we infer the convergence $(F^{X_d})^{-1}(r) \rightarrow \ln r$, $d \rightarrow \infty$, for any $r \in (0, 1)$. In particular, we have

$$\lim_{d \rightarrow \infty} (F^{X_d})^{-1}(1 - \varepsilon^2) = \ln(1 - \varepsilon^2) \quad \text{for all } \varepsilon \in (0, 1).$$

It is true that $(F^{X_d})^{-1}(1 - \varepsilon^2) = -d \tilde{\varphi}(|\ln \bar{\lambda}^{X_d}(\varepsilon)|)$ for any $\varepsilon \in (0, 1)$ and $d \in \mathbb{N}$. Indeed, by (20), we have

$$\begin{aligned} d \tilde{\varphi}(|\ln \bar{\lambda}^{X_d}(\varepsilon)|) &= d \tilde{\varphi}\left(\inf\left\{x \in \mathbb{R} : \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_d} \mathbf{1}(|\ln \bar{\lambda}_k^{X_d}| \leq x) \geq 1 - \varepsilon^2\right\}\right) \\ &= d \tilde{\varphi}\left(\inf\left\{x \in \mathbb{R} : \sum_{k=1}^{\infty} \bar{\lambda}_k^{X_d} \mathbf{1}\left(d \tilde{\varphi}(|\ln \bar{\lambda}_k^{X_d}|) \geq d \tilde{\varphi}(x)\right) \geq 1 - \varepsilon^2\right\}\right) \\ &= -\inf\{y \in \mathbb{R} : F^{X_d}(y) \geq 1 - \varepsilon^2\} \\ &= -(F^{X_d})^{-1}(1 - \varepsilon^2). \end{aligned}$$

Hence for any $\varepsilon \in (0, 1)$ we obtain $\lim_{d \rightarrow \infty} d \tilde{\varphi}(|\ln \bar{\lambda}^{X_d}(\varepsilon)|) = -\ln(1 - \varepsilon^2)$. By the strict decay of $\tilde{\varphi}$, for any $\varepsilon \in (0, 1)$ $|\ln \bar{\lambda}^{X_d}(\varepsilon)| \rightarrow \infty$ as $d \rightarrow \infty$. The convergence (61) follows from (62).

Next, we prove (60). From (14) and strict decay of $\tilde{\varphi}$ it follows that

$$\varliminf_{d \rightarrow \infty} d \tilde{\varphi}(\ln n^{X_d}(\varepsilon)) \geq \lim_{d \rightarrow \infty} d \tilde{\varphi}(|\ln \bar{\lambda}^{X_d}(\varepsilon)|) = -\ln(1 - \varepsilon^2).$$

Fix arbitrary $h > 0$ and $c_h \in (1, 1/\varepsilon)$ such that $\ln(1 - c_h^2 \varepsilon^2)/\ln(1 - \varepsilon^2) < e^h$. Using (15) with $\varepsilon_1 = \varepsilon$ and $\varepsilon_2 = c_h \varepsilon$, $|\ln \bar{\lambda}^{X_d}(\varepsilon)| \rightarrow \infty$, $d \rightarrow \infty$, and slow variation of $\tilde{\varphi}$, we obtain:

$$\begin{aligned} \overline{\lim}_{d \rightarrow \infty} d \tilde{\varphi}(\ln n^{X_d}(\varepsilon)) &\leq \lim_{d \rightarrow \infty} d \tilde{\varphi}(|\ln \bar{\lambda}^{X_d}(\varepsilon)| + \ln((c_h^2 - 1)\varepsilon^2/2)) \\ &= \lim_{d \rightarrow \infty} d \tilde{\varphi}(|\ln \bar{\lambda}^{X_d}(\varepsilon)|) \\ &= -\ln(1 - c_h^2 \varepsilon^2) \\ &\leq -\ln(1 - \varepsilon^2)e^h. \end{aligned}$$

Hence $\lim_{d \rightarrow \infty} d \tilde{\varphi}(\ln n^{X_d}(\varepsilon)) = -\ln(1 - \varepsilon^2)$ for any $\varepsilon \in (0, 1)$. The equivalence (62) yields (60). \square

Corollary 1 *For any $\varepsilon \in (0, 1)$ $(\ln n^{X_d}(\varepsilon))_{d \in \mathbb{N}}$ is rapidly varying sequence, i. e.*

$$\lim_{d \rightarrow \infty} \frac{\ln n^{X_{\lfloor cd \rfloor}}(\varepsilon)}{\ln n^{X_d}(\varepsilon)} = \infty \quad \text{for all } c > 1,$$

where $\lfloor \cdot \rfloor$ is a floor function.

Thus here growth of $n^{X_d}(\varepsilon)$ is extremely fast as $d \rightarrow \infty$ (it can be double exponential, $\exp\{\exp\{\cdot\}\}$, see an example in the next subsection).

To prove this corollary we suppose, contrary to our claim that for some $c > 1$ there exists a subsequence $(d_k)_{k \in \mathbb{N}}$ such that

$$\lim_{k \rightarrow \infty} \frac{\ln n^{X_{\lfloor cd_k \rfloor}}(\varepsilon)}{\ln n^{X_{d_k}}(\varepsilon)} = p > 0. \tag{63}$$

On the one hand, by Theorem 9, we obtain

$$\lim_{k \rightarrow \infty} \left(\lfloor cd_k \rfloor \varphi(\ln n^{X_{\lfloor cd_k \rfloor}}(\varepsilon)) \right) = -\ln(1 - \varepsilon^2).$$

On the other hand, from slow variation of φ and (63) it follows

$$\lim_{k \rightarrow \infty} \lfloor cd_k \rfloor \varphi(\ln n^{X_{\lfloor cd_k \rfloor}}(\varepsilon)) = \lim_{k \rightarrow \infty} \lfloor cd_k \rfloor \varphi(\ln n^{X_{d_k}}(\varepsilon)) = -c \ln(1 - \varepsilon^2).$$

This leads to a contradiction.

5.2 Applications

Let us consider a sequence of tensor degrees $X_d = X^{\otimes d}$, $d \in \mathbb{N}$. Suppose that eigenvalues of K^X has the following asymptotics:

$$\frac{\lambda_k^X}{\Lambda^X} \sim \frac{\beta}{k^p (\ln k)^{1+r}}, \quad k \rightarrow \infty, \quad (64)$$

where $\beta > 0$ and numbers p and r must satisfy $\sum_{k=1}^{\infty} \lambda_k^X < \infty$, i.e. $p > 1$, $r \in \mathbb{R}$ or $p = 1$, $r > 0$.

The following assertion is a direct corollary of Theorem 7.

Proposition 8 *Under the assumption (64) with $\beta > 0$, $p > 1$, $r \in \mathbb{R}$, or $p = 1$, $r > 2$, we have*

$$\ln n^{X_d}(\varepsilon) = E^X d + \Phi^{-1}(1 - \varepsilon^2) \sigma^X d^{1/2} + o(d^{1/2}), \quad d \rightarrow \infty, \quad \text{for all } \varepsilon \in (0, 1),$$

where E^X and σ^X are defined by (53) and (54), respectively.

In order to consider the remaining cases we need the following auxiliary lemma.

Lemma 2 *Under the assumption (64) with $p = 1$, $r > 0$, we have*

$$\sum_{k=1}^{\infty} \bar{\lambda}_k^X \mathbf{1}(\bar{\lambda}_k^X < e^{-x}) \sim \frac{\beta}{r} \cdot x^{-r}, \quad x \rightarrow \infty.$$

Proof of Lemma 2. First, we get

$$\sum_{k=n}^{\infty} \bar{\lambda}_k^X \sim \sum_{k=n}^{\infty} \frac{\beta}{k (\ln k)^{1+r}} \sim \int_n^{\infty} \frac{\beta dt}{t (\ln t)^{1+r}} = \frac{\beta}{r} \cdot (\ln n)^{-r}, \quad n \rightarrow \infty. \quad (65)$$

Set $k_x := \min\{k \in \mathbb{N} : \bar{\lambda}_k^X < e^{-x}\}$. Using Lemmas 5 and 6 (see Appendix), we find

$$k_x \sim e^x ((\ln)^{1+r} / \beta)^{\#}(e^x) \sim \frac{\beta e^x}{x^{1+r}}, \quad x \rightarrow \infty, \quad (66)$$

and, in particular, $\ln k_x \sim x$, $x \rightarrow \infty$. From this and (65) we obtain the required asymptotics. \square

Proposition 9 *Under the assumption (64) with $\beta > 0$, $p = 1$, and $r = 2$, for all $\varepsilon \in (0, 1)$ we have*

$$\ln n^{X_d}(\varepsilon) = E^X d + \Phi^{-1}(1 - \varepsilon^2) (\beta/2)^{1/2} (d \ln d)^{1/2} + o((d \ln d)^{1/2}), \quad d \rightarrow \infty,$$

where E^X is defined by (53).

Proof of Proposition 9. On account of Lemma 2 for $r = 2$, the required asymptotics is obtained by Theorem 6 with $\varphi(x) \sim \beta/2$, $x \rightarrow \infty$, and $a_d = E^X d$, $b_d := d^{1/2} \sqrt{2} (1/\sqrt{\hat{\varphi}})^\#(d^{1/2})$, where $\hat{\varphi}$ is defined by (56). We only need to find asymptotics of the sequence $(b_d)_{d \in \mathbb{N}}$. It easily seen that

$$\hat{\varphi}(x) = \int_0^x \frac{\hat{\varphi}(t)}{t} dt \sim \int_1^x \frac{\beta}{2t} dt \sim \frac{\beta \ln x}{2}, \quad x \rightarrow \infty.$$

Hence, from Lemma 6 (see Appendix) we have

$$(1/\sqrt{\hat{\varphi}})^\#(d^{1/2}) \sim \hat{\varphi}(d^{1/2})^{1/2} \sim ((\beta/4) \ln d)^{1/2}, \quad d \rightarrow \infty.$$

This yields $b_d \sim (\beta/2)^{1/2} (d \ln d)^{1/2}$, $d \rightarrow \infty$. \square

Proposition 10 *Under the assumption (64) with $\beta > 0$, $p = 1$, and $r \in (0, 2)$, we have*

$$\ln n^{X_d}(\varepsilon) = a_d + S_{r,1}^{-1}(1 - \varepsilon^2) (\beta/r)^{1/r} d^{1/r} + o(d^{1/r}), \quad d \rightarrow \infty, \quad \text{for all } \varepsilon \in (0, 1),$$

where $a'_d := 0$ for $r \in (0, 1)$, $a'_d := dE^X$ for $r \in (1, 2)$, and

$$a_d := d \sum_{k=1}^{\infty} |\ln \bar{\lambda}_k^X| \bar{\lambda}_k^X \mathbf{1}(\bar{\lambda}_k^X > e^{-\beta d}) + (1 - \mathcal{C})\beta d, \quad \text{for } r = 1. \quad (67)$$

Here \mathcal{C} is the Euler constant and E^X is defined by (53). In the case $r = 1$ $a_d \sim \beta d \ln \ln d$, $d \rightarrow \infty$.

Proof of Proposition 10. According to Lemma 2 with given $r \in (0, 2)$, we can use Theorem 6 with $\alpha = r$ and $\varphi(x) \sim \beta/r$, $x \rightarrow \infty$. Since, by Lemma 6, we have

$$d^{1/r} ((1/\varphi)^{1/r})^\#(d^{1/r}) \sim d^{1/r} \varphi(d^{1/r})^{1/r} \sim (\beta/r)^{1/r} d^{1/r}, \quad d \rightarrow \infty,$$

on account of Lemma 3, numbers b_d can be choosen as follows $b_d := (\beta/r)^{1/r} d^{1/r}$, $d \in \mathbb{N}$. The expressions for a_d are directly obtained from Theorem 6. In particular, for the case $r = 1$ we have the formula (67). On account of (66), we get $k_{\beta d} := \min\{k \in \mathbb{N} : \bar{\lambda}_k^X < e^{-\beta d}\} \sim e^{\beta d}/(\beta d^2)$. Accordingly, using assumptions on $\bar{\lambda}_k^X$, $k \in \mathbb{N}$, it follows that

$$\sum_{k=1}^{\infty} |\ln \bar{\lambda}_k^X| \bar{\lambda}_k^X \mathbf{1}(\bar{\lambda}_k^X \geq e^{-\beta d}) \sim \sum_{k=1}^{k_{\beta d}} \frac{\beta}{k \ln k} \sim \ln \ln k_{\beta d} \sim \beta \ln \ln d, \quad d \rightarrow \infty.$$

This gives $a_d \sim \beta d \ln \ln d$, $d \rightarrow \infty$, for the case $r = 1$. \square

In order to have full expansion of a_d (up to $o(d)$) in Proposition 10 for the case $r = 1$, we need know more information about asymptotic behaviour of $\bar{\lambda}_k^X$, $k \in \mathbb{N}$. We omit these details, which reduce to routine calculations.

Now we consider an example of applying Theorem 9.

Proposition 11 *Suppose that*

$$\frac{\lambda_k^X}{\Lambda^X} \sim \frac{\beta}{k(\ln k)(\ln \ln k)^{1+s}}, \quad k \rightarrow \infty,$$

with $\beta > 0$ and $s > 0$. Then

$$\ln \ln n^{X_d}(\varepsilon) \sim \left(\frac{\beta d}{s |\ln(1 - \varepsilon^2)|} \right)^{1/s}, \quad d \rightarrow \infty, \quad \text{for all } \varepsilon \in (0, 1).$$

Proof of Proposition 11. First, we get

$$\sum_{k=n}^{\infty} \bar{\lambda}_k^X \sim \sum_{k=n}^{\infty} \frac{\beta}{k(\ln k)(\ln \ln k)^{1+s}} \sim \int_n^{\infty} \frac{\beta dt}{t \ln t (\ln \ln t)^{1+s}} = \frac{\beta}{s} \cdot (\ln \ln n)^{-s}, \quad n \rightarrow \infty.$$

Set $k_x := \min\{k \in \mathbb{N} : \bar{\lambda}_k^X < e^{-x}\}$. Using Lemma 5 (see Appendix) and Lemma 6, we find

$$k_x \sim e^x ((\ln)(\ln \ln)^{1+s}/\beta)^{\#}(e^x) \sim \frac{\beta e^x}{x(\ln x)^{1+s}}, \quad x \rightarrow \infty,$$

that gives $\ln \ln k_x \sim \ln x$, $x \rightarrow \infty$. From this we have

$$\sum_{k=1}^{\infty} \bar{\lambda}_k^X \mathbf{1}(\bar{\lambda}_k^X < e^{-x}) = \sum_{k=k_x}^{\infty} \bar{\lambda}_k^X \sim \frac{\beta}{s} \cdot (\ln x)^{-s}, \quad x \rightarrow \infty.$$

The required asymptotics conclude from Theorem 9. \square

6 Appendix

Here we recall the definitions and basic properties of self-decomposable and stable probability distributions. Also we provide known limit theorems of weak convergence to these distributions. The facts and notions are used in the previous sections. For more detailed study see [9], [13], and [30].

6.1 Self-decomposable distributions

Definition 1 A distribution function F is called *self-decomposable* if for any $\alpha > 1$ there exists a distribution function F_α such that $F(x) = \int_{\mathbb{R}} F_\alpha(x-y) dF(\alpha y)$ for all $x \in \mathbb{R}$.

Self-decomposable distribution functions are also called *distributions functions of class \mathbf{L}* (or *\mathbf{L} functions* for short). Every self-decomposable distribution function F is infinitely divisible and thus its characteristic function $f(t) = \int_{\mathbb{R}} e^{itx} dF(x)$, $t \in \mathbb{R}$, uniquely admits *Lévy canonical representation*:

$$f(t) = \exp \left\{ i\gamma t - \frac{\sigma^2 t^2}{2} + \int_{|x|>0} \left(e^{itx} - 1 - \frac{itx}{1+x^2} \right) dL(x) \right\}, \quad t \in \mathbb{R}, \quad (68)$$

where $\gamma \in \mathbb{R}$, $\sigma^2 \geq 0$, *Lévy spectral function* $L: \mathbb{R} \setminus \{0\} \rightarrow \mathbb{R}$ is non-decreasing on the intervals $(-\infty, 0)$ and $(0, \infty)$ and it satisfies the conditions

$$\begin{aligned} \lim_{x \rightarrow -\infty} L(x) &= \lim_{x \rightarrow \infty} L(x) = 0, \\ \int_{0 < |x| < \tau} x^2 dL(x) &< \infty \quad \text{for all } \tau > 0. \end{aligned}$$

Only for \mathbf{L} functions its Lévy spectral functions L are continuous and have one-sided derivatives on $\mathbb{R}/\{0\}$, where the functions $x \mapsto xL'(x)$ are non-increasing on the intervals $(-\infty, 0)$ and $(0, \infty)$ (here $L'(x)$ is a left or right derivative of L at x).

It is known fact that any \mathbf{L} function is unimodal (see [44] and also [42]). Any non-degenerate \mathbf{L} function is absolutely continuous (see [33], [43] and [45]).

Remark 8 If F is a non-degenerate \mathbf{L} function, then it strictly increases on $(\text{lex} F, \text{rex} F)$.

This fact follows from theorem of W. N. Hudson and H. G. Tucker (see [12]), which states that a set of zeroes for density of arbitrary absolutely continuous infinitely divisible distribution function either has Lebesgue measure zero or almost surely coincides with some infinite interval.

For example, a distribution function D_β with the following characteristic function

$$f_{D_\beta}(t) = \exp \left\{ \beta \int_0^1 \frac{e^{itx} - 1}{x} dx \right\}, \quad \beta > 0, \quad t \in \mathbb{R},$$

is self-decomposable. It has canonical Lévy representation with the triplet

$$\gamma = \int_0^1 \frac{\beta dx}{1+x^2} = \frac{\beta\pi}{4}, \quad \sigma^2 = 0, \quad L(x) = \beta(\ln x) \mathbf{1}(x \in (0, 1]).$$

Density ρ_β of the distribution function D_β equals zero on $(-\infty, 0]$ and it satisfies the following equation (see [11], [42], and [43])

$$x\rho'_\beta(x) = (\beta - 1)\rho_\beta(x) - \beta\rho_\beta(x - 1), \quad x > 0,$$

with the initial condition

$$\rho_\beta(x) = \frac{e^{-\beta\mathcal{C}}}{\Gamma(\beta)} \cdot x^{\beta-1}, \quad 0 < x \leq 1,$$

where \mathcal{C} is the Euler constant.

The function ρ , defined by the equation $\rho(x) = e^{\mathcal{C}}\rho_1(x)$, $x \in \mathbb{R}$, is called *Dickman function*. It occupies an important place in number theory (see [34]). According to this, the distribution corresponding to D_1 is called the *Dickman distribution* and the distributions corresponding to D_β , $\beta > 0$, are called *convolution powers of the Dickman distribution* (see [11]).

Self-decomposable distributions are of interest because of the following theorem (see [30], p. 101).

Theorem 10 Let $(A_n)_{n \in \mathbb{N}}$ be a sequence and $(B_n)_{n \in \mathbb{N}}$ be a positive sequence. Let $(Y_j)_{j \in \mathbb{N}}$ be a sequence of independent random variables satisfying the following condition

$$\lim_{n \rightarrow \infty} \max_{j=1, \dots, n} \mathbb{P}(|Y_j| > xB_n) = 0 \quad \text{for all } x > 0. \quad (69)$$

If the ditribution functions of the following sums

$$\frac{\sum_{j=1}^n Y_j - A_n}{B_n}, \quad n \in \mathbb{N}, \quad (70)$$

weakly converge to a non-degenerate distribution function F , then $F \in \mathbf{L}$ and $B_n \rightarrow \infty$, $B_{n+1}/B_n \rightarrow 1$ as $n \rightarrow \infty$.

Let us formulate necessary and sufficient conditions for convergence of distribution functions of (70) to a given \mathbf{L} function (see [9], p. 124).

Theorem 11 *Let $(A_n)_{n \in \mathbb{N}}$ be a sequence and $(B_n)_{n \in \mathbb{N}}$ be a positive sequence. Let $(Y_j)_{j \in \mathbb{N}}$ be a sequence of independent random variables satisfying (69). Let F be an \mathbf{L} function with triplet (γ, σ^2, L) in Lévy canonical representation (68). For the distribution functions of the sums (70) to converge weakly to F , it is necessary and sufficient to satisfy*

$$\begin{aligned}
(A_1) \quad & \lim_{n \rightarrow \infty} \sum_{j=1}^n \mathbb{P}(Y_j > xB_n) = -L(x) \quad \text{for all } x > 0; \\
(A_2) \quad & \lim_{n \rightarrow \infty} \sum_{j=1}^n \mathbb{P}(Y_j \leq xB_n) = L(x) \quad \text{for all } x < 0; \\
(B) \quad & \lim_{n \rightarrow \infty} \frac{1}{B_n} \left(\sum_{j=1}^n \mathbb{E} \left[Y_j \mathbf{1}(|Y_j| \leq \tau B_n) \right] - A_n \right) = \\
& \quad = \gamma + \int_{0 < |x| < \tau} \frac{x^3 \, dL(x)}{1+x^2} - \int_{|x| \geq \tau} \frac{x \, dL(x)}{1+x^2} \quad \text{for all } \tau > 0; \\
(C) \quad & \lim_{\tau \rightarrow 0} \lim_{n \rightarrow \infty} \frac{1}{B_n^2} \sum_{j=1}^n \mathbf{Var} \left[Y_j \mathbf{1}(|Y_j| \leq \tau B_n) \right] = \lim_{\tau \rightarrow 0} \overline{\lim}_{n \rightarrow \infty} \frac{1}{B_n^2} \sum_{j=1}^n \mathbf{Var} \left[Y_j \mathbf{1}(|Y_j| \leq \tau B_n) \right] = \sigma^2.
\end{aligned}$$

The convergence of (70) still holds if we replace $(B_n)_{n \in \mathbb{N}}$ with any equivalent sequence. It follows from the next general lemma (see [30] p. 21).

Lemma 3 *Let $(c_n)_{n \in \mathbb{N}}$ be a sequence, $r = (r_n)_{n \in \mathbb{N}}$ be a positive sequence. Suppose that distribution functions F_n weakly converge to a non-degenerate distribution function F . Then the following assertions hold:*

- (i) *If $F_n(r_n x + c_n)$ weakly converge to non-degenerate distribution function H , then $H(x) = F(rx + c)$, where $c = \lim_{n \rightarrow \infty} c_n$ and $r = \lim_{n \rightarrow \infty} r_n$.*
- (ii) *If $c = \lim_{n \rightarrow \infty} c_n$ and $r = \lim_{n \rightarrow \infty} r_n$, then $F_n(r_n x + c_n)$ weakly converge to $F(rx + c)$.*

6.2 Stable distributions

Definition 2 *A distribution function F is called stable if for any $a_1 > 0$ and $a_2 > 0$ there exist $a > 0$ and b such that $F(ax + b) = \int_{\mathbb{R}} F(a_1(x - y)) \, dF(a_2 y)$ for all $x \in \mathbb{R}$.*

It is well known that every stable distribution function is self-decomposable and thus infinitely divisible. The class of stable distributions consists degenerate distributions, normal distributions and non-normal α -stable distributions, $\alpha \in (0, 2)$.

The characteristic function of normal distribution function Φ_{μ, σ^2} (with mean $\mu \in \mathbb{R}$ and variance $\sigma^2 > 0$) has representation (68) with the triplet $(\mu, \sigma^2, 0)$, where Lévy spectral function is identically equals zero. Standard normal distribution function $\Phi_{0,1}$ is denoted by Φ , i.e.

$$\Phi(x) = \Phi_{\mu, \sigma^2}(\mu + \sigma x). \quad (71)$$

We will denote by $S_{\alpha,\rho,\beta,\mu}$ the stable distribution functions, which are non-degenerate and non-normal. The parametrization is chosen according to so called *A*-form of representation of their characteristic functions (see [4] p. 10):

$$f_{S_{\alpha,\rho,\beta,\mu}}(t) := \exp\left\{i\mu t - \rho\kappa_\alpha|t|^\alpha(1 - i\beta\operatorname{sign}(t)\omega(t,\alpha))\right\},$$

where $\alpha \in (0, 2)$, $\rho > 0$, $\beta \in [-1, 1]$, $\mu \in \mathbb{R}$, and

$$\kappa_\alpha := \begin{cases} \frac{\Gamma(2-\alpha)\cos(\pi\alpha/2)}{1-\alpha}, & \alpha \in (0, 2), \alpha \neq 1, \\ \pi/2, & \alpha = 1; \end{cases}$$

$$\omega(t, \alpha) := \begin{cases} \tan(\pi\alpha/2), & \alpha \in (0, 2), \alpha \neq 1, \\ -(2/\pi)\ln|t|, & \alpha = 1. \end{cases}$$

We will call $S_{\alpha,\beta} := S_{\alpha,1,\beta,0}$, $\alpha \in (0, 2)$, $\beta \in [-1, 1]$, the *standard α -stable distribution functions*, which can be obtained by the formula

$$S_{\alpha,\beta}(x) = S_{\alpha,\rho,\beta,\mu}(\mu_1 + \rho^{1/\alpha}x),$$

where $\mu_1 := \mu + (\beta\rho\ln\rho)\mathbf{1}(\alpha = 1)$.

The characteristic function of $S_{\alpha,\rho,\beta,\mu}$ has representation (68) with the triplet $(\gamma, 0, L)$. Here

$$L(x) = \frac{\rho(1-\beta)}{2|x|^\alpha}\mathbf{1}(x < 0) - \frac{\rho(1+\beta)}{2|x|^\alpha}\mathbf{1}(x > 0), \quad x \in \mathbb{R} \setminus \{0\}, \quad (72)$$

and $\gamma = \mu + \alpha\beta\rho I_\alpha$, where

$$I_\alpha = \int_0^\infty \left(\frac{1}{x^\alpha(1+x^2)} - \frac{\sin x}{x^2}\mathbf{1}(\alpha = 1) - \frac{1}{x^\alpha}\mathbf{1}(1 < \alpha < 2) \right) dx.$$

This integral can be explicitly computed. By formula **3.781** 1. in [8], we find

$$I_1 = \int_0^\infty \left(\frac{1}{x(1+x^2)} - \frac{\sin x}{x^2} \right) dx = \mathcal{C} - 1,$$

where $\mathcal{C} \approx 0,5772$ is the Euler constant. By formula **3.241** 2. in [8], we have

$$\int_0^\infty \frac{x^p dx}{1+x^2} = \frac{\pi}{2\cos(\pi p/2)}, \quad p \in (-1, 1).$$

Hence we obtain

$$I_\alpha = \int_0^\infty \frac{x^{-\alpha} dx}{1+x^2} = \frac{\pi}{2\cos(\pi\alpha/2)} \quad \text{for } \alpha \in (0, 1),$$

and also

$$I_\alpha = \int_0^\infty \left(\frac{1}{x^\alpha(1+x^2)} - \frac{1}{x^\alpha} \right) dx = - \int_0^\infty \frac{x^{2-\alpha} dx}{1+x^2} = \frac{\pi}{2 \cos(\pi\alpha/2)} \quad \text{for } \alpha \in (1, 2).$$

The analytic properties of stable distributions was in detail described in the monograph [46], the related limit theorems — in [4]. The general reviews can be found in classic monographs [6], [9], [13], [22], and [32] but it should be take into account the paper [10].

The fundamental role of stable distributions is explained by the following theorem.

Theorem 12 *The set of distribution functions that are weak limits of the distribution functions of centered and normalized sums (70) with independent and identically distributed random variables Y_j , $j \in \mathbb{N}$, coincides with the set of stable distribution functions.*

The criterion of convergence to standard normal distribution function has the following form (see Theorem 2.6.2 in [13], but there is a typos in the formulation: “if” should be changed to “iff”).

Theorem 13 *Let $(Y_j)_{j \in \mathbb{N}}$ be a sequence of non-degenerate independent and identically distributed random variables. There exist a sequence $(A_n)_{n \in \mathbb{N}}$ and a strictly positive sequence $(B_n)_{n \in \mathbb{N}}$ such that the distribution functions of (70) weakly convergence as $n \rightarrow \infty$ to Φ iff the following conditions hold:*

- (i) $\text{Var } Y_1 < \infty$,
- (ii) $\mathbb{P}(|Y_1| > x) = x^{-2}\varphi(x)$,

where φ is a SVF³. For A_n , $n \in \mathbb{N}$, it can be set $A_n = n \mathbb{E} Y_1$, the B_n , $n \in \mathbb{N}$, can be taken from $n \text{Var} (Y_1 \mathbb{1}(|Y_1| < B_n)) \sim B_n^2$, $n \rightarrow \infty$. In case (i) we can set $B_n = \sqrt{n \text{Var } Y_1}$.

The criterion of convergence to standard α -stable distribution functions has the following form (see [32] p. 50, [41] p. 114).

Theorem 14 *Let $(Y_j)_{j \in \mathbb{N}}$ be a sequence of non-degenerate independent and identically distributed random variables. There exist a sequence $(A_n)_{n \in \mathbb{N}}$ and a strictly positive sequence $(B_n)_{n \in \mathbb{N}}$ such that the distribution functions of (70) weakly convergence as $n \rightarrow \infty$ to $S_{\alpha, \beta}$, $\alpha \in (0, 2)$, $\beta \in [-1, 1]$, iff both*

- 1) $\mathbb{P}(|Y_1| > x) = x^{-\alpha}\varphi(x)$,
- 2) $\lim_{x \rightarrow \infty} \frac{\mathbb{P}(Y_1 > x)}{\mathbb{P}(|Y_1| > x)} = \frac{1 + \beta}{2}$,

where φ is a SVF. The B_n , $n \in \mathbb{N}$, can be taken from $\lim_{n \rightarrow \infty} n \mathbb{P}(|Y_1| > B_n) = 1$. The A_n , $n \in \mathbb{N}$, can be chosen as follows: $A_n = 0$ for $\alpha \in (1, 2)$, $A_n = n \mathbb{E} Y_1$ for $\alpha \in (1, 2)$, and

$$A_n = n \mathbb{E} (Y_1 \mathbb{1}(Y_1 \leq B_n)) + \beta(1 - \mathcal{C})B_n \quad \text{for } \alpha = 1.$$

³slowly varying function on infinity, see Subsection 6.4

For the case $\alpha = 1$ the expression of the A_n , $n \in \mathbb{N}$, is often omitted in the literature or has often a difficult form. Let us explain the possibility of the choice of A_n , $n \in \mathbb{N}$, in Theorem 14. Indeed, by the condition (B) of Theorem 11 for $\tau = 1$ and on account of above expressions of γ and L for $\mu = 0$, $\rho = 1$, and $\alpha = 1$, the constants A_n , $n \in \mathbb{N}$, must be satisfy

$$\lim_{n \rightarrow \infty} \frac{1}{B_n} \left(\sum_{j=1}^n \mathbb{E} Y_j \mathbf{1}(Y_j \leq B_n) - A_n \right) = \beta(\mathcal{C} - 1) + \beta \left(\int_0^1 \frac{x \, dx}{1+x^2} - \int_1^\infty \frac{dx}{(1+x^2)x} \right).$$

The last integrals are equal, therefore, we have the required formula for A_n , $n \in \mathbb{N}$.

6.3 The Darling theorem

Suppose that common distribution function of independent and identically distributed random variables has a slowly varying summary tail, i.e. $\mathbb{P}(|Y_1| > x) = \varphi(x)$, where φ is some SVF that $\varphi(x) \rightarrow 0$ as $x \rightarrow \infty$. In this case the distribution functions of the sums

$$\frac{\sum_{j=1}^{n_k} Y_j - A_k}{B_k}, \quad k \in \mathbb{N},$$

can not have non-degenerate weak limit for any $(A_k)_{k \in \mathbb{N}}$, $(B_k)_{k \in \mathbb{N}}$, and $(n_k)_{k \in \mathbb{N}}$ (see [6] p. 320). However, here *the Darling theorem* holds (see [5]). We formulate it in the form, which was obtained by S. V. Nagaev and V. I. Vakhtel in the paper [24]. Also we restrict ourselves to the case of non-negative random variables.

Theorem 15 *Let $(Y_j)_{j \in \mathbb{N}}$ be a sequence of non-degenerate independent and identically distributed non-negative random variables such that $\mathbb{P}(Y_1 > x) = \varphi(x)$ with some SVF φ . Then we have*

$$\lim_{n \rightarrow \infty} \mathbb{P} \left(n \tilde{\varphi} \left(\sum_{i=1}^n Y_i \right) \leq x \right) = 1 - e^{-x} \quad \text{for all } x > 0,$$

where $\tilde{\varphi}(y) := 1 - \mathbb{E} e^{-Y_1/y}$, $y > 0$.

In this theorem the function $\tilde{\varphi}$ is continuous and strictly decreasing on $(0, \infty)$. Also, according to tauberian theorem (see [6], p. 447, formula (5.22)), we have $\tilde{\varphi}(x) \sim \varphi(x)$, $x \rightarrow \infty$.

6.4 Some facts from theory of regular variation

Let φ be a positive measurable function, defined on some $[T, \infty)$ and satisfying $\varphi(cx)/\varphi(x) \rightarrow 1$, $x \rightarrow \infty$ for any $c > 0$. Then φ is said to be a *slowly varying* at infinity (SVF for short). Here we provide some useful lemmas concerning asymptotic inversion and conjugation of such functions.

Lemma 4 *Suppose that φ is a SVF. Then there exists a SVF $\varphi^\#$, unique up to asymptotic equivalence, that satisfies*

1. $\lim_{x \rightarrow \infty} \varphi(x) \varphi^\#(x \varphi(x)) = 1$,
2. $\lim_{x \rightarrow \infty} \varphi^\#(x) \varphi(x \varphi^\#(x)) = 1$,

3. $\varphi^{\#\#}(x) \sim \varphi(x)$, $x \rightarrow \infty$.

The SVF $\varphi^\#$ is called *de Bruijn conjugation* of φ . The importance of $\varphi^\#$ is explained by the following lemma (see [2] p. 28–29).

Lemma 5 *Let $f(x) \sim x^\tau \varphi(x)^\tau$, $x \rightarrow \infty$ with $\tau > 0$, φ is a SVF. There exists the function g that $f(g(x)) \sim g(f(x)) \sim x$ and $g(x) \sim x^{1/\tau} \varphi^\#(x^{1/\tau})$, $x \rightarrow \infty$, where $\varphi^\#$ is the de Bruijn conjugation for φ . Here g is determined uniquely up to asymptotic equivalence, and one version of g is $f^{-1}(x) = \inf\{y \in \mathbb{R} : f(y) > x\}$.*

For many cases $\varphi^\#$ may be asymptotically expressed in terms of φ itself ([2] p. 78–79).

Lemma 6 *Let φ be a SVF with the de Bruijn conjugation $\varphi^\#$. If $\varphi(x) \sim \varphi(x/\varphi(x))$, $x \rightarrow \infty$, then we have $\varphi^\#(x) \sim 1/\varphi(x)$, $x \rightarrow \infty$.*

Let us mention one more useful assertion (see [2] p. 26–27).

Lemma 7 *Let φ be a SVF that locally integrable on $[T, \infty)$ for some $T \in \mathbb{R}$. Let function $\hat{\varphi}$ define by $\hat{\varphi}(x) := \int_0^x (\varphi(t)/t) dt$. Then $\hat{\varphi}$ is a SVF and $\hat{\varphi}(x)/\varphi(x) \rightarrow \infty$, $x \rightarrow \infty$.*

Acknowledgments

The author wishes to thank the Chebyshev Laboratory of St. Petesburg State University for providing excellent working conditions. Also the author wishes to express his gratitude to Professor M. A. Lifshits for many helpful suggestions and comments during the preparation of the paper.

References

- [1] R. J. Adler, J. Taylor, *Random Fields and Geometry*, Springer, New York, 2007.
- [2] N. H. Bingham, C. M. Goldie, J. L. Teugels, *Regular Variation*, Camb. Univ. Press, Cambridge, 1987.
- [3] C.-H. Chang, C.-W. Ha, *The Green functions of some boundary value problems via the Bernoulli and Euler polynomials*, Arch. Math. (Basel), **76** (2001), no. 5, 360–365.
- [4] G. Christoph, W. Wolf, *Convergence Theorems with a Stable Limit Law*, Math. Research 70, Akad. Verlag, Berlin, 1992.
- [5] D. A. Darling, *The Influence of the maximum term in the addition of independent random variables*, Trans. Amer. Math. Soc., **73** (1952), no. 1, 95–107.
- [6] W. Feller, *Introduction to Probability Theory and Its Applications*, vol. 2, Wiley, New York, 1971.
- [7] F. Gao, J. Hanning, F. Torcaso, *Integrated Brownian motions and exact L_2 -small balls*, Ann. Probab., **31** (2003), no. 3, 1320–1337.
- [8] I. S. Gradshteyn, I. M. Ryzhik, *Tables of Integrals, Series and Products*, Elsevier, Burlington, 2007.

- [9] B. V. Gnedenko, A. N. Kolmogorov, *Limit Distributions for Sums of Independent Random Variables*, Addison-Wesley, Cambridge, 1954.
- [10] P. Hall, *A comedy of errors: the canonical form for a stable characteristic function*, Bull. London Math. Soc., **13** (1981), no. 1, 23–27.
- [11] D. Hensley, *The convolution powers of the Dickman function*, J. London Math. Soc., **33** (1986), no. 3, 395–406.
- [12] W. N. Hudson, H. G. Tucker, *On admissible translates of infinitely divisible distributions*, Z. Wahrsch. Verw. Gebiete, **31** (1975), 65–72.
- [13] I. A. Ibragimov, Yu. V. Linnik, *Independent and Stationary Sequences of Random Variables*, Wolters-Noordhoff, Groningen, 1971.
- [14] A. Karol, A. Nazarov, Ya. Nikitin, *Small ball probabilities for Gaussian random fields and tensor products of compact operators*, Trans. Amer. Math. Soc., **360** (2008), no. 3, 1443–1474.
- [15] M. A. Lifshits, *Lectures on Gaussian Processes*, Springer, New York, 2012.
- [16] M. A. Lifshits, A. Papageorgiou, H. Woźniakowski, *Average case tractability of non-homogeneous tensor product problems*, J. Complexity, **28** (2012), no. 5–6, 539–561.
- [17] M. A. Lifshits, A. Papageorgiou, H. Woźniakowski, *Tractability of multi-parametric Euler and Wiener integrated processes*, Probab. Math. Stat., **32** (2012), no. 1, 131–165.
- [18] M. A. Lifshits, E. V. Tulyakova, *Curse of dimensionality in approximation of random fields*, Probab. Math. Stat., **26** (2006), no. 1, 97–112.
- [19] M. A. Lifshits, M. Zani, *Approximation complexity of additive random fields*, J. Complexity, **24** (2008), no. 3, 362–379.
- [20] M. A. Lifshits, M. Zani, *Approximation of additive random fields based on standard information: average case and probabilistic settings*, [arXiv:1402.5489](https://arxiv.org/abs/1402.5489)
- [21] Yu. V. Linnik, I. V. Ostrovskii, *Decomposition of Random Variables and Vectors*, Transl. Math. Monog., Vol. 48, AMS, Providence, Rhode Island, 1977.
- [22] M. Loève, *Probability Theory I*, vol. 1, Grad. Text Math. 45, Springer-Verlag, New-York, 1977.
- [23] B. M. Makarov, M. G. Goluzina, A. A. Lodkin, A. N. Podkorytov, *Selected Problems in Real Analysis*, Transl. Math. Monog., vol. 107, AMS, Providence, Rhode Island, 1992.
- [24] S. V. Nagaev, V. I. Vachtel, *On sums of independent random variables without power moments*, Siberian Math. J., **49** (2008), no. 6, 1091–1100.
- [25] A. I. Nazarov, Ya. Yu. Nikitin, *Exact L_2 -small ball behavior of integrated Gaussian processes and spectral asymptotics of boundary value problems*, Probab. Theory Relat. Fields, **129** (2004), no. 4, 469–494.
- [26] E. Novak, I. H. Sloan, J. F. Traub, H. Woźniakowski, *Essays on the Complexity of Continuous Problems*, EMS, Zürich, 2009.

- [27] E. Novak, H. Woźniakowski, *Tractability of Multivariate Problems. Volume I: Linear Information*, EMS Tracts Math. 6, EMS, Zürich, 2008.
- [28] E. Novak, H. Woźniakowski, *Tractability of Multivariate Problems. Volume II: Standard Information for Functionals*, EMS Tracts Math. 12, EMS, Zürich, 2010.
- [29] E. Novak, H. Woźniakowski, *Tractability of Multivariate Problems. Volume III: Standard Information for Operators*, EMS Tracts Math. 18, EMS, Zürich, 2012.
- [30] V. V. Petrov, *Limit Theorems of Probability Theory: Sequences of Independent Random Variables*, Oxford Stud. Prob. 4, Clarendon Press, Oxford, 1995.
- [31] K. Ritter, *Average-case Analysis of Numerical Problems*, Lecture Notes in Math. No. 1733, Springer, Berlin, 2000.
- [32] G. Samorodnitsky, M. S. Taqqu, *Stable Non-Gaussian Random Processes: Stochastic Models with Infinite Variance*, Chapman & Hall, London, 1994.
- [33] K.-I. Sato, M. Yamazato, *On Distribution function of class L* , Z. Wahrsch. Verw. Gebiete, **43** (1978), 273–308.
- [34] G. Tenenbaum, *Introduction to Analytic and Probabilistic Number Theory*, Camb. Univ. Press, Cambridge, 1995.
- [35] J. F. Traub, G. W. Wasilkowski, H. Woźniakowski, *Information, Uncertainty, Complexity*, Addison-Wesley, Reading MA, 1983.
- [36] J. F. Traub, G. W. Wasilkowski, H. Woźniakowski, *Information-Based Complexity*, Academic Press, New York, 1988.
- [37] J. F. Traub, H. Woźniakowski, *A general theory of optimal algorithms*, Academic Press, New York, 1980.
- [38] J. F. Traub, A. G. Werschulz, *Complexity and Information*, Camb. Univ. Press, Cambridge, 1998.
- [39] A. W. van der Vaart, *Asymptotic Statistics*, Camb. Univ. Press, Cambridge, 1998.
- [40] G. W. Wasilkowski, H. Woźniakowski, *Average case optimal algorithms in Hilbert spaces*, J. Approx. Theory, **47** (1986), 17–25.
- [41] W. Whitt, *Stochastic-Process Limits: An Introduction to Stochastic-Process Limits and Their Application to Queues*, Springer, New York, 2002.
- [42] S. J. Wolfe, *On the unimodality of L functions*, Ann. Math. Statist., **42** (1971), no. 3, 912–918.
- [43] S. J. Wolfe, *On the continuity properties of L functions*, Ann. Math. Statist., **42** (1971), no. 6, 2064–2073.
- [44] M. Yamazato, *Unimodality of infinitely divisible distribution functions of class L* , Ann. Probab., **6** (1978), no. 4, 523–531.

- [45] V. M. Zolotarev, *The analytic structure of infinitely divisible laws of class L*, Litovsk. Mat. Sb., **3** (1963), 123–140.
- [46] V. M. Zolotarev, *One-dimensional Stable Distributions*, Transl. Math. Monog., vol. 65, AMS, Providence, Rhode Island, 1986.

Keywords and phrases: multivariate problems, tensor product-type random elements, approximation, average case approximation complexity, asymptotic analysis, tractability.

CHEBYSHEV LABORATORY, ST. PETERSBURG STATE UNIVERSITY, 14TH LINE, 29B, ST. PETERSBURG, 199178 RUSSIA

E-mail address: alexeykhartov@gmail.com